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## UNIVERSITY OF CALIFORNIA

Los Angeles

Sexual orientation differences in distal and proximal risk factors of suicide

A dissertation submitted in partial satisfaction of the requirements for the degree Doctor of

Philosophy in Epidemiology

by

Kirsty Amy Clark

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#### ABSTRACT OF THE DISSERTATION

Sexual orientation differences in distal and proximal risk factors of suicide

by

Kirsty Amy Clark

Doctor of Philosophy in Epidemiology

University of California, Los Angeles, 2019

Professor Susan D. Cochran, Chair

The Surgeon General classifies sexual minorities (i.e., lesbians, gay men, bisexuals) as a high-risk population for suicide. This three-paper dissertation sought to contribute to an understanding of suicide risk among sexual minorities by investigating sexual orientation differences in both distal (i.e., stigmatization) and proximal (e.g., access to firearms) cofactors of suicide. Specifically, the studies investigated: (Study 1) the impact of neighborhood-level structural stigma on changes in psychological distress among sexual minorities following a statewide vote to ban same-sex marriage, (Study 2) sexual orientation differences in gun ownership and gun safety policy beliefs, and (Study 3) sexual orientation differences in lethal method used in suicide. Data were drawn from three unique datasets, including the California Quality of Life Survey (Study 1; N=7,421), the General Social Survey (Study 2; N=4,554), and

the National Violent Death Reporting System (Study 3; N=59,075). Statistical analyses were conducted in R and SAS 9.4 with analytic methods across studies including univariate and bivariate analyses, quasi-experimental methods, multivariate logistic regression adjusted for confounding, and complex survey weighting incorporating propensity score methods. Results from Study 1 demonstrated that sexual minorities residing in neighborhoods where a majority of constituents voted in favor of a same-sex marriage ban had a statistically significant mean psychological distress reduction in the time period immediately after the vote ( $\beta$ =-4.06, 95% CI: -7.38; -0.73), with post-hoc analyses demonstrating that increased social support might have partially influenced the direction of these unexpected findings. Study 2 findings demonstrated that gay/bisexual men were more likely to endorse a gun safety law (aOR=3.24, 95% CI: 1.79 -5.88) and less likely to report guns in the household (aOR=0.34, 95% CI: 0.18 - 0.65) than heterosexual men. Among women reporting a household gun, lesbian/bisexual women were more likely to be the personal gun owner (aOR=3.97, 95% CI: 1.43 - 11.03) than heterosexual women. Results from Study 3 showed that sexual minorities who died by suicide had reduced odds of death by firearm (aOR=0.62, 95% CI: 0.51 - 0.75) and increased odds of death by hanging (aOR=1.36, 95% CI: 1.14 - 1.63) and drug/poison ingestion (aOR=1.30, 95% CI: 1.04 -1.62) as compared to the reference group of suicides; however, these patterns varied by gender. Findings from this dissertation demonstrate that sexual orientation differentially impacts distal (i.e., psychological responses to structural stigma) and proximal (i.e., access to firearms, use of lethal method in suicide) risk factors for suicide. Results from the research reported here can inform public health and clinical interventions to reduce suicide risk among sexual minorities.

The dissertation of Kirsty Amy Clark is approved.

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2019

### **DEDICATION**

I dedicate this dissertation to my father, Ross Henry Clark, who believed in me, nurtured my intellectual curiosities, and continually told me he was proud of me. His mantra was, "if you think you can, you can."

I thought I could, and I did, Dad. This is for you.

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**Clark, K.,** Blosnich, J.R., Coulter, R.W.S., Bamwine, P., Bossarte, R.M. & Cochran, S.D. (In Press) Sexual orientation differences in gun ownership and beliefs about gun safety policy, General Social Survey 2010-2016. *Violence and Gender*.

- Clark, K., Blosnich, J.R., Haas, A.P. & Cochran, S.D. (2019) Estimate of Lesbian, Gay, Bisexual and Transgender Youth Suicide is Inflated. *Journal of Adolescent Health*, 64(6), 810.
- Reback, C. J., **Clark, K.,** Fletcher, J. B., & Holloway, I. W. (2019). A Multilevel Analysis of Social Network Characteristics and Technology Use on HIV Risk and Protective Behaviors Among Transgender Women. *AIDS and Behavior*, 1-15.
- White Hughto, J. & Clark, K. (2019) Designing a Transgender Health Training for Correctional Healthcare Providers: A Feasibility Study. *The Prison Journal*.
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### Chapter 1

### Suicide, Stigma, and Sexual Minorities: A Brief Introduction

Reducing suicide represents a critical intervention point for public health. Suicide<sup>a</sup> is a leading cause of preventable mortality globally. Worldwide, suicide results in approximately 800,000 deaths per year, accounting for 1.4% of all deaths.<sup>1</sup> Annually, over 55,000 deaths in the U.S. are due to violence-related injury, almost two-thirds of which are attributable to suicide.<sup>2</sup> Increasingly, it has become a public health imperative to reduce suicide through public health and clinical interventions targeting groups disproportionately impacted by self-inflicted violent death.<sup>3-5</sup> For this dissertation, I executed three related studies focused on the precursors of suicide. My purpose is to inform suicide prevention efforts among sexual minorities (e.g., lesbians, gay men, bisexuals), a population vulnerable to suicide.<sup>6,7</sup> As a foundation for the work, I herein provide a brief epidemiological and theoretical background of suicide and suicide risk generally and among sexual minorities; propose an updated conceptual model of suicide risk among sexual minorities; and finally, outline the later dissertation chapter topics.

#### **Suicide and Suicide Risk Factors**

Suicide is the tenth leading cause of death in the U.S., and among 15-19 year olds, it is the second leading cause of death after unintentional injury.<sup>8</sup> A 2018 report from the Center for Disease Control and Prevention (CDC) demonstrated that since 1999, suicide rates in the U.S. have risen nearly 30%.<sup>9</sup> In 2016, approximately 1.3 million U.S. adults attempted suicide (0.5%)

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<sup>&</sup>lt;sup>a</sup> <u>Suicide</u> is defined by the National Institute of Mental Health as, "death caused by self-directed injurious behavior with intent to die as a result of the behavior." In this dissertation, I use the terms "suicide", "death by suicide", and "suicide mortality". The terms "killed him/herself," "completed suicide," "successful suicide," and "committed suicide" are avoided given negative connotations and best practices and preferred terminology from suicide prevention organizations (e.g. National Institute of Mental Health, Beyond Blue, Suicide Prevention Lifeline, and Suicide Prevention Resource Center).

of the population).<sup>10</sup> Studies have demonstrated that 8.8% - 13% of suicide attempts are lethal; young women, blacks, and adolescents have higher rates of suicide attempt while men, whites, and the elderly have higher rates of suicide mortality.<sup>11,12</sup>

Suicide does not have well-recognized risk mechanisms that can invariably predict its occurrence. It is perhaps best understood as an adverse outcome generated by a complex web of social, psychological, behavioral and environmental risk factors, none of which are essential for causation. 13 Suicide prevention is, therefore, challenging because specific, sufficient causes of suicide are elusive. Since Emile Durkheim's seminal 1897 study of religious, regional, and demographic differences in suicide, <sup>14</sup> numerous studies have attempted to characterize specific risks of suicide attempt and mortality, including psychiatric disorders and substance use, 15-17 family history, <sup>18,19</sup> stigmatization, <sup>20</sup> and even contagion (i.e., one suicide prompting another). <sup>21,22</sup> Studies investigating the epidemiology of suicide demonstrate that distal risk factors for suicide include psychological, developmental, biological, and familial risks, while proximal risk factors include situational contexts unique to the suicide attempt, including whether or not the individual had access to a firearm.<sup>23,24</sup> Distal and proximal risk factors are not mutually exclusive, and it is indeed their co-occurrence that exacerbates suicide risk.<sup>23</sup> Studies differentiating risk factors for suicide attempt versus mortality have demonstrated that risk factors for suicide mortality (as opposed to attempt) include a prior suicide attempt, severe psychiatric illness, and use of a firearm.<sup>25-27</sup>

#### **Interpersonal Theory of Suicide (ITS)**

Relatively few theoretical models have been proposed to explain the etiology of suicide mortality. The Interpersonal Theory of Suicide (ITS)<sup>28</sup> emerged in 2005 to fill this gap, asserting that *perceived burdensomeness* (i.e., the perception that one's existence burdens family, friends,

and/or society) and *thwarted belongingness* (i.e., disconnectedness from others) foster suicidal desire. However, the ITS asserts that it is only in the context of an *acquired capability* for lethal self-injury that suicide results. This acquired capability for suicide refers to the ability to overcome the evolutionary reflex for self-preservation required to end one's life.<sup>29</sup> **Figure 1.1** provides a model of the Interpersonal Theory of Suicide.

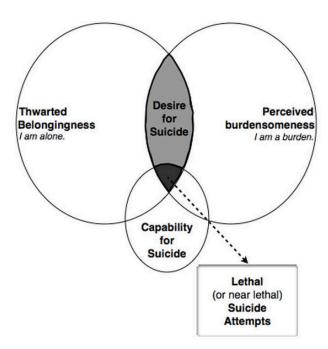


Figure 1.1. Interpersonal Theory of Suicide (from Van Orden et al., 2015)

Support for the ITS has been demonstrated across numerous diverse populations and national and international samples.<sup>30</sup> One set of studies empirically testing the ITS demonstrated three key findings supporting the interactions between the concepts outlined in the ITS: (1) the interaction of perceived burdensomeness and thwarted belongingness predicted current suicidal ideation; (2) individuals with a greater number of past suicide attempts demonstrated higher levels of acquired capability; and, (3) an interaction between acquired capability and perceived

burdensomeness was associated with clinical risk of suicidal behavior.<sup>31,32</sup> Recent tests have further shown that painful and provocative experiences (e.g., substance abuse, exposure to self-harm, witnessing or experiencing trauma or violence) are positively associated with suicidal intent and increased capability for suicide.<sup>33</sup>

#### What about Lethal Method? A Shortcoming of the ITS

Despite being the primary theoretical model for conceptualizing pathways to lethal or nearlethal suicide attempts, the ITS fails to include the principal predictor of whether or not a suicide attempt will ultimately result in mortality: the method used in the attempt. 12 While an individual may indeed possess the desire and capability to end their life, having access to a lethal method (e.g., firearm, high bridge) is the final gatekeeper for suicide death. 34 Suicide mortality has been linked to impulsivity and limited planning, <sup>35,36</sup> especially among adolescents; <sup>37</sup> thus, the immediate accessibility of a lethal method is paramount to the lethality of any suicide attempt. 34,38 Means restriction interventions have been effective in reducing firearms suicides in the U.S. and Austria, <sup>39-42</sup> carbon monoxide and barbiturate suicides in Denmark. <sup>43,44</sup> gas poisoning suicides in England and Wales, 45,46 alcohol-related suicides in Slovenia and Russia, 47,48 and jumping-related suicides across numerous international high-bridge sites. 49-51 But despite the documented success of restricting lethal means in suicide prevention, this approach has historically been overlooked by policymakers presumably given its environmental nature. 52,53 Thus, a 2012 call-to-action from researchers at Harvard School of Public Health emphasized the need to develop and implement public health and clinical interventions focused on restricting access to highly-lethal methods (e.g., firearms) to reduce suicide mortality.<sup>54</sup> Relatedly, theoretical models attempting to show pathways to suicide mortality should be cognizant to include lethal method as the final pathway through which suicide occurs.

In population-based studies investigating case-fatality rates in suicide attempts, researchers consider firearm, hanging, and drowning highly lethal suicide means as they are associated with greater than 80% fatality rates. Less lethal means include poisoning with drugs. Here case fatality estimates range from 2% to 14% 11,12 The lethal method used in a suicide attempt not only varies by regional accessibility to the method (e.g., the disproportionately high firearms suicides in the U.S. as compared to other countries), but also varies by individual status characteristics. For example, the "gender paradox" in suicide – that women are more likely to attempt suicide, but men are more likely to die by suicide – can be partially attributed to the fact that men are more likely than women to use a highly lethal method of suicide in their attempt. 55-57 Older age is also positively associated with using highly lethal suicide means. 12,58 Studies have also demonstrated race-related differences in suicide method, including one study showing that black men were more likely to use highly lethal suicide means as compared to white men. <sup>59,60</sup> Understanding differences in suicide method based on key status characteristics can not only augment our understanding of the epidemiology of suicide but can also aid in targeting interventions to groups disproportionately vulnerable to death by suicide.

### **Sexual minority status**

Sexual orientation has been defined in varying ways for over 160 years, <sup>61</sup> encompassing the many facets of sexual attraction, sexual behavior, and self-identity. <sup>62</sup> As of 2019, the American Psychological Association defines sexual orientation as, "an often enduring pattern of emotional, romantic and/or sexual attractions of men to women or women to men (heterosexual), of women to women or men to men (homosexual, gay, lesbian), or by men or women to both sexes (bisexual)." 'Sexual minority' is an umbrella term for individuals who are not heterosexual, or part of the 'sexual majority'. Whether or not a person of same-sex attraction chooses to identify

openly as a sexual minority, or "come out," hinges on complex social and cultural influences including the supportiveness of family, involvement in religion, and community acceptance. 63,64

Sexual orientation identity development is hypothesized to occur across the life course in key stages spanning pre-puberty to adulthood. 63,65 There are numerous stage-sequential models theorizing the processes involved in sexual minority identity development. 66-69 These models vary in terminology and theoretical underpinnings; however, they generally follow a similar trajectory of development milestones. As a specific example of one of the earliest model's stages, Troiden (1988) defined the four stages of homosexual identity development as: (1) Sensitization in childhood, characterized by the perceptions of being different from same-sex peers, often including gender-discordant interests and behavior; (2) Identity confusion in adolescence, characterized by the thought that one might be homosexual, which is dissonant with prior self-images, and often leads to anxiety, guilt, and social isolation; (3) Identity assumption in which one's identity is established and shared with 'safe' others including other homosexuals; and (4) Commitment, which is characterized by living openly as a sexual minority and integrating sexual minority identity into all facets of life, including entering a same-sex committed sexual and/or romantic relationship. Homosexual identity development models have been updated more recently to highlight that while the ordering of "coming-out" milestones might be similar across individuals, the timing of these events across the life course varies, spanning childhood to older adulthood. 70,71

#### **Stigma & Sexual Minorities**

Stigma is defined as a "mark or label" (attribute, characteristic) that is devalued in a particular social context, 72 and reduces the stigmatized "from a whole and usual person to a tainted, discounted one." Being a sexual minority is considered a stigmatized identity, 74 and

ample evidence demonstrates that lesbians, gay men, and bisexual people experience stigmatization across the life course. As well, research suggests that some sexual minority populations experience the additive effects of multiple stigmatized identities; for example, the sexism, racism, and homophobia experienced by black sexual minority women has been dubbed a "triple jeopardy," exacerbating negative health outcomes. 81,82

Herek (2009) defines "sexual stigma" as a specific type of stigma "attached to any nonheterosexual behavior, identity, relationship, or community." Sexual stigma functions at two levels: 4 (1) structural-level stigma, which refers to the "policies of private and governmental institutions that intentionally restrict the opportunities" of stigmatized groups, including same-sex marriage bans, anti-gay employment discrimination policies, and laws banning adoption by same-sex couples; and (2) interpersonal or individual-level stigma, which refers to negative attitudes towards sexual minorities by heterosexuals, or negative attitudes towards oneself (i.e., internalized stigma) among sexual minorities.

Minority stress theory (MST) posits that chronic stress, generated from actual and anticipated experiences of stigmatization, contributes to disproportionate mental health burden among sexual minorities, such as increased rates of psychological distress including depression, anxiety and substance use disorder, which independently predict suicidality.<sup>75,77,86-92</sup>

#### **Psychological Mediation Framework (PMF)**

The Psychological Mediation Framework (PMF) builds on the MST to provide a theorydriven model explicating the intrapsychic and interpersonal pathways at play in the relationship between stigma-related stressors and psychological distress among sexual minorities. Indeed, the PMF posits that psychological distress and stigma-related stressors are mediated through three key intermediaries: emotion dysregulation, social/interpersonal issues, and maladaptive cognitive states. 93 **Figure 1.2** depicts the PMF.

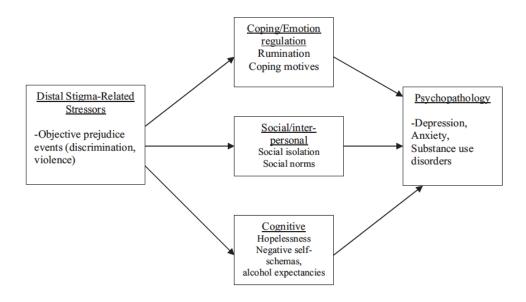


Figure 1.2. Psychological Mediation Framework (from Hatzenbuehler, 2009)

The PMF specifically posits that stress generated by sexual orientation stigma (e.g., hate crimes, victimization, discrimination)<sup>76</sup> leaves sexual minorities more vulnerable than heterosexuals to psychological processes that cause psychopathology. <sup>93</sup> The PMF elaborates on the MST, which does not explicitly posit specific intra- or interpersonal mediators to explain the jump in the pathway from stress to psychopathology. In this regard, the PMF identifies the specific pathways through which stigma-related stressors generate psychological distress, including depression, anxiety and, importantly, *suicidality*.

#### Suicide risk among sexual minorities

For over 30 years, research has consistently demonstrated that sexual minorities have higher rates of stress-related psychological distress as compared to heterosexuals, including elevated rates of mental disorders and suicidality. <sup>6,90,91,94,95</sup> Indeed, the Surgeon General classifies sexual

minorities as a high-risk population for suicide. Gompared to heterosexuals, sexual minorities are at increased risk of suicide attempt, the most powerful clinical predictor of suicide mortality. Population-based research from the U.S. finds that sexual minorities are 2-7 times more likely to attempt suicide than similar heterosexuals. As well, men and women in same-sex partnerships demonstrate disproportionate risk for suicide mortality. While numerous studies document elevated risk of suicidality among sexual minorities as compared to heterosexuals, Res is understood about sexual orientation differences in proximal precursors of suicide mortality (i.e., circumstances immediately preceding death). Such knowledge might facilitate development of effective intervention and prevention strategies.

#### **Proposed Conceptual Model of Suicide among Sexual Minorities**

There is currently no widely-used conceptual model of suicide mortality specific to sexual minorities. This is warranted for two primary reasons: (1) sexual minorities demonstrate elevated risk of suicide ideation, attempt and mortality as compared to heterosexuals; and, (2) factors that exacerbate suicide risk among sexual minorities are likely different, in some cases, than those factors impacting heterosexual suicide risk (e.g., sexual minorities' disproportionate exposure to stigmatization). Thus, I herein propose a conceptual model of suicide among sexual minorities that guides the trajectory of my dissertation research. **Figure 1.3** melds two aforementioned theories, the Psychological Mediation Framework and the Interpersonal Theory of Suicide, along with method of suicide, into one integrated conceptual model of suicide among sexual minorities.

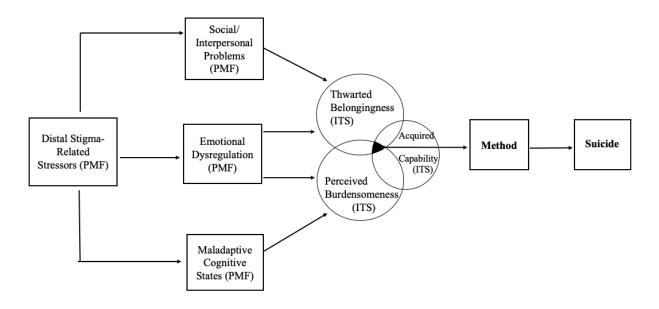


Figure 1.3. An adapted theoretical model of suicide among sexual minorities integrating the Interpersonal Theory of Suicide (ITS) and Psychological Mediation Theory (PMF)

The PMF provides explicit mechanisms whereby stigma leaves sexual minorities vulnerable to: (1) social/interpersonal problems, including low social support and exclusion, that contribute to *thwarted belongingness* (ITS), (2) emotional dysregulation, including impulsivity, negative coping strategies, and hypervigilance, that contribute to *thwarted belongingness* and *perceived burdensomeness* (ITS) and, (3) maladaptive cognitive states, including hopelessness and low self-esteem that contribute to *perceived burdensomeness* (ITS).<sup>28</sup> In the context of acquired capability, and through the 'final gatekeeper' of method of suicide, suicide results. Empirical research and experimental evidence provide support for the proposed conceptual model. Social and interpersonal problems, including social exclusion and ostracism, drive social isolation and thwart the need to belong. <sup>101,102</sup> Emotionally dysregulated individuals evidence higher levels of both perceived burdensomeness and thwarted belongingness. <sup>103</sup> And maladaptive cognitive states (i.e., hopelessness and negative self-schemas) exacerbate the perception of being a burden on others. <sup>104</sup> Sexual minorities report elevated levels of thwarted

belongingness, emotional dysregulation, and perceived burdensomeness as compared to similar heterosexuals. 90,105-110

My dissertation does not attempt to test all pathways outlined in this conceptual model; rather I focus on two disparate pieces of this model. First, I investigate the association between a specific, distal stigma-related stressor and psychological distress among sexual minorities. Then, I explore the relationship between sexual orientation and access to firearms, the most lethal suicide method. Finally, I evaluate sexual orientation differences in method of suicide. This dissertation attempts to strengthen our understanding of key risk factors to suicide especially relevant for sexual minorities: stigmatization and differential access to and use of lethal method in suicide. This project has the potential to identify clinical and public health points of intervention in suicide prevention efforts among sexual minorities.

### Chapter 2

Measuring the impact of neighborhood-level structural stigma on psychological distress among sexual minorities: A natural experiment drawing on California's 2008 Proposition 8 Vote

### Background

Stigma is considered a fundamental cause of population health inequalities.<sup>111</sup> Among sexual minorities (i.e., lesbians, gay men, and bisexuals [LGB]), a preponderance of evidence suggests that stigma drives the population's disproportionate risk for adverse mental health outcomes (e.g., depression, anxiety, substance use) as compared to similar heterosexuals.<sup>74,78,93,111,112</sup> Stigma experiences can include both *interpersonal-level stigmatization* (e.g., violence, rejection, bullying) and *structural-level stigmatization* (e.g., discriminatory laws and policies, anti-LGB social norms).

#### Interpersonal-level stigma

The majority of stigma-related research on sexual minorities has focused on the adverse mental health impact of interpersonal stigma (e.g., enacted discrimination, prejudice). Minority stress theory (MST) posits that chronic stress, generated from actual and anticipated experiences of stigmatization, contributes to disproportionate mental health burden among sexual minorities, including increased rates of depression, anxiety and substance use disorder, which then predict suicidality. 75,77,86-92 Experiences of stigma and resultant psychological distress have been linked to suicidal ideation 113 and suicide attempt 114-116 across numerous studies. Further, research has demonstrated that individuals with histories of painful and provocative experiences (e.g., violence, sexual abuse) might be at elevated risk of death by suicide because painful experiences increase a suicidal individual's pain tolerance and reduce the fear of death. 29,117,118 As compared to heterosexuals, sexual minorities report disproportionate experiences of interpersonal

discrimination, abuse, and violent victimization across the life course; <sup>79,91,107,114,119,120</sup> thus, it is pertinent that interpersonal stigma and violent victimization are addressed when developing effective suicide prevention strategies.

### Structural-level stigma

Recent research has also investigated the mechanisms by which structural-level stigma impacts psychological distress and suicidality among sexual minorities. Indeed, recent quasiexperimental studies have attempted to demonstrate the impact of statewide same-sex marriage bans and statewide laws permitting denial of services to same-sex couples on sexual minority mental health. A quasi-natural experiment found that the passing of a state same-sex marriage law in Massachusetts in 2003 led to a reduction in medical and mental healthcare visits in the 12 months following the enactment among sexual minority men who entered into treatment just prior to the law's passing. 121 This study was conducted with a small clinic sample and did not utilize a control group, significantly detracting from its generalizability and the study's attempt at causal inference. In the past two years, Raifman et al. have capitalized on policy implementation, large-scale survey data, and utilized causal inference methodology to investigate the causal impact of structural stigma on sexual minority mental health. 122,123 One such study used data from 762,678 students in the Youth Risk Behavior Surveillance System (YRBSS) from 1999 to 2015 to compare adolescent suicide attempts before and after the implementation of policies permitting same-sex marriage in 32 states, using 15 states without such policies as a control group. 123 Findings demonstrated a 7% relative reduction in suicide attempts due to same-sex marriage implementation, with this association concentrated among sexual minorities. A second study used data from 109,089 participants in the Behavioral Risk Factor Surveillance System (BRFSS) to demonstrate that implementation of laws permitting denial of services to same-sex

couples in three states was associated with a 46% increase in sexual minority adults experiencing psychological distress. 122 These studies demonstrate that structural stigma operating in the form of a discriminatory state-level policy might negatively impact sexual minority mental health. However, these studies provide only the state-level context of the impact of structural stigma; indeed, neither study attempted to explore how a statewide policy might differentially impact sexual minority individuals living in more- or less-stigmatizing geographic regions of the state. State-level vs. neighborhood-level structural stigma

Understanding how laws and policies affect sexual minority mental health at the state level provides a bird's-eye view of the impact of structural stigma on mental health; however, there is less evidence demonstrating how state-level stigmatization might differentially affect sexual minority psychological distress depending on an individual's neighborhood-level characteristics. Neighborhood-level characteristics might be potent indicators of health outcomes among sexual minorities; indeed, research has demonstrated that some neighborhood characteristics are associated with sexual minority health. Several studies have investigated stigma-related neighborhood-level characteristics including the proportion of households with same-sex couples, the proportion of lesbian, gay, bisexual or transgender (LGBT) hate crimes, and the proportion of Republican voters to show that these factors influence health outcomes including mental health, substance use, and sexual risk behaviors among sexual minorities. 124-126 As well, a spatial analysis from Columbus, Ohio demonstrated that sexual minorities were concentrated within neighborhoods with other sexual minorities, and among gay men, geographic patterns of housing were related to diversity, openness, and amenities in the neighborhood. 127 Indeed, the term "gayborhood" has been used historically to refer to gay neighborhoods: geographical safe

spaces in which sexual minorities are concentrated, and in which there are clusters of gay-owned or gay-friendly businesses, meeting spaces, and organizations. <sup>128</sup>

Thus, while the aforementioned studies investigating structural stigma at the state level highlight the possible impact of statewide policies on sexual minority mental health, an in-depth investigation into how neighborhood-level characteristics might buffer or exacerbate psychological distress in relation to a state-level event (e.g., same-sex marriage ban) can offer insight into how neighborhood councils and organizations might tackle stigmatization at a local level. It is plausible that sexual minorities living in neighborhoods with climates supportive to homosexuality – even in states that pass stigmatizing state-level policies – might be more resilient against the harmful mental health effects of state-level structural stigmatization compared to sexual minorities facing both state-level and neighborhood-level stigmatization. One study from Australia used voting data from a non-binding referendum on same-sex marriage to investigate the influence of geographic variation in community-level structural stigma on mental health, life satisfaction, and overall health among sexual minorities. 129 This study showed that sexual minorities living in local areas with a higher percentage of constituents voting against same-sex marriage had worse health outcomes. But this finding was mediated by social support such that social support buffered these negative effects. However, this study utilized crosssectional data at a single time point after the referendum had already occurred; thus, it is unknown how the vote itself might have impacted sexual minority mental health (i.e., researchers were unable to map trends in mental health before and after the vote).

In sum, while prior studies have provided some evidence for the causal impact of stigmatizing state- and federal-level policy on sexual minority mental health, <sup>122,123</sup> and others have utilized cross-sectional data to demonstrate an association between neighborhood- and

community-level structural stigma and sexual minority mental health, <sup>126,129</sup> our current understanding of how neighborhood-level characteristics might differentially influence sexual minority mental health in the wake of a stigmatizing state-level policy is quite limited.

Thus, the current study exploited a natural experiment to understand how a stigmatizing state-level vote might differentially impact psychological distress among sexual minorities depending on neighborhood variation in voting behavior. Specifically, I investigated the differential impact of a statewide same-sex marriage ban (i.e., California's Proposition 8) on self-reported recent psychological distress among sexual minorities using a quasi-experimental design. Below I introduce California's political landscape and provide a brief summary of California's Proposition 8.

### California Political Geography & Proposition 8

California's Political Landscape

California's political landscape is complex. While California decidedly votes Democratic in presidential elections, it has strong regional variations in voting behavior and ideological public opinions, commonly associated with a "coastal-inland divide" (with the coast typically voting Democratic and the inland regions typically voting Republican). Figure 2.1 shows a map from the Public Policy Institute of California highlighting this geographic divide with data from the 2008 Presidential Election by county in California. Thus, while California might have a reputation for espousing primarily liberal views, there is a strong regional variation that is tied to voting behavior and opinions on social issues.

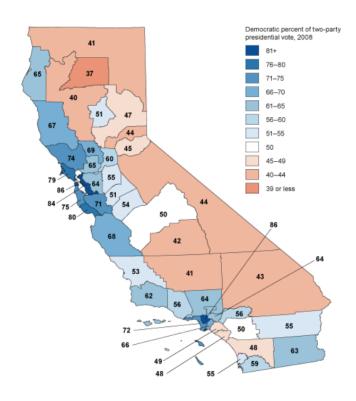


Figure 2.1. Map of Democratic percent of the two-party presidential vote by county in California, 2008

### California's Proposition 8

On June 16<sup>th</sup>, 2008, same-sex marriage was legalized in California following the Supreme Court of California ruling in *In re Marriage Cases* which found that excluding same-sex couples from marrying was a violation of the California constitution.<sup>131</sup> This resulted in populace backlash and on November 4<sup>th</sup>, 2008, Proposition 8 (hereafter Prop 8), passed in the 2008 California statewide election. Prop 8 amended the California constitution to state that "only marriage between a man and a woman is valid or recognized in California," effectively banning same-sex marriage.<sup>131</sup> The proposition passed by a close margin: 7,001,084 (52.3%) voted in favor of Prop 8 and 6,401,482 (47.7%) voted against it. **Figure 2.2** provides a visualization of the geographic variation of voting behavior for Prop 8 by California counties. Note that a "Yes"

vote on Prop 8 was associated with a vote to ban same-sex marriage, while a "No" vote was associated with a rejection of constitutional wording changes to ban same-sex marriage.

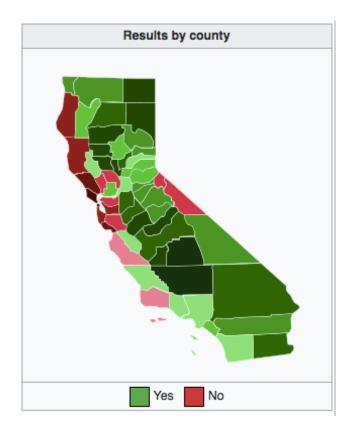


Figure 2.2. Distribution of Prop 8 Vote by California county

The Prop 8 vote was highly publicized on social media and in the news. <sup>132,133</sup> A cross-sectional convenience study conducted in the days leading up to the passing of Prop 8 found that sexual minorities reported negative emotions including feeling "upset", "nervous", and "angry." Of the study sample (N= 354), approximately 75% answered "7" (*A great deal*) out of a 7-point Likert scale when asked, "how much time have you spent thinking about Proposition 8?" Because Prop 8 revoked an existing right to marry, sexual minorities in California might have experienced greater psychological distress to the marriage ban than sexual minorities in other

states.<sup>84,134</sup> Thus, there is solid evidence that Prop 8 was a salient matter, certainly to sexual minorities in California, and when passed served to enact an event consistent with structural stigmatization.

Study Purpose & Hypothesis

The current study sought to understand how neighborhood-level structural stigma (as indexed by the proportion of one's neighborhood that voted in favor of Prop 8, a same-sex marriage ban) impacted sexual minority psychological distress in the wake of a stigmatizing 'shock' (i.e., Prop 8). To investigate this, I used data from multiple waves of a representative statewide survey in California from time periods before, during, and after the Prop 8 vote. Given the robust literature demonstrating the strong harmful associations between structural stigma and sexual minority mental health, I hypothesized that sexual minorities residing in neighborhoods with a higher level of structural stigma would demonstrate increased levels of psychological distress in the time period immediately after the same-sex marriage ban was enacted as compared to sexual minorities living in neighborhoods with a lower level of structural stigma. This study expands on prior literature by utilizing a quasi-experimental design to offer insight into how a stigmatizing 'shock' in the population might differentially impact psychological distress depending on the stigma-related characteristics of one's local environment.

#### Methods

#### Data

Data were drawn from three waves of the California Quality of Life Survey (Cal-QOL). Each Cal-QOL is a follow-up to one of three waves of the California Health Interview Survey (CHIS). The CHIS is the largest state health survey in the United States, and at the time employed random-digit-dial telephone procedures to interview Californians about their health and

healthcare needs. From 2001-2012, the CHIS was conducted bi-annually, each time collecting data from between 42,000 and 55,000 Californian households per wave. Across years, the overall CHIS response rate was consistent with other random-digit-dial telephone interviews of the same time period, including the California Behavioral Risk Factor Surveillance System Survey. 135-137

Follow-back surveys like the Cal-QOL use the obtained sample in the original survey sample (i.e., CHIS) as the sampling frame to procure a subsample of interest. <sup>136</sup> In the CHIS, respondents aged 18 to 70 years were asked their sexual orientation identity and gender of recent sexual partners. Eligibility for the Cal-QOL included those CHIS respondents who were willing to be re-contacted for additional health surveys and completed an English or Spanish version of the CHIS interview. From this eligible sample, the Cal-QOL oversampled for sexual minority status. Cal-QOL data were collected through structured interviews with trained interviewers using Computer Assisted Telephone Interviews (CATI). Interviews were conducted in English or Spanish.

Cal-QOL-I participants were interviewed between October 2004 and February 2005 with a 56% response rate; Cal-QOL-II participants were interviewed between August 2008 and January 2009 with a 57% response rate; and Cal-QOL-III participants were interviewed between January 2012 and February 2013 with a 58% response rate.

Participants were considered a sexual minority if they identified as lesbian, gay or bisexual at the time of the Cal-QOL interview, or if they reported lifetime histories of same-sex sexual behavior. Because sexual minorities were oversampled, the final Cal-QOL dataset includes 7,536 participants, including 2,021 sexual minorities and 5,515 heterosexuals (Cal-QOL-I n = 2,272 collected between October 2004 and February 2005; Cal-QOL-II n = 2,815 collected between

August 2008 and January 2009; Cal-QOL-III n = 2,449 collected between January 2012 and February 2013).

#### **Variables**

Outcome Variable

The dependent variable for this study is the Kessler Psychological Distress Scale (K-10) score. This 10-item scale measures psychological distress in the past 30 days (see **Appendix 2.A** for the full scale). Scores range from 0 to 40. Previous research linking K-10 scores to clinical measures of mental health morbidity suggest the following cutoff scores: no/low levels of psychological distress (0-9), mild psychological distress (10-14), moderate psychological distress (15-19), and severe psychological distress (20-40). Preliminary analyses of the K-10 distribution in the total Cal-QOL sample revealed that the average weighted K-10 score was 4.96 (Std Err: 0.09) and the weighted distribution of K-10 classification was as follows: 84.1% met criteria for no/low distress, 8.5% met criteria for mild distress, 3.6% met criteria for moderate distress, and 3.7% met criteria for severe distress. Because the K-10 measures current or recent (prior 30 days) psychological distress, it is an ideal measure with which to capture changes in population distress over time and in response to a shock event. As well, the K-10 has been used as a valid measure of psychological distress in numerous studies investigating sexual minority mental health. 137,141,142

Treatment Variable: Neighborhood-level Structural Stigma

The treatment variable for this study is a binary measure of structural stigma generated from precinct-level voting data and individual-level geographic variation by neighborhood.

Participants were allocated into 'treated' (i.e., resides in a neighborhood where a majority voted in favor of same-sex marriage ban) or 'untreated' (i.e., resides in a neighborhood where a

majority voted against same-sex marriage ban) groups. How this variable was created is described below.

Zip Code Tabulation (ZCTA) overview. The Cal-QOL obtained information on respondents' zipcode and majority cross-streets during the course of the interview. This was converted to ZCTAs. Created by the U.S. Census Bureau, ZCTAs are generalized area representations of U.S. Postal Service (USPS) service areas.<sup>143</sup> In most cases, a ZCTA is the same as a zip code for a geographic area, though this is not always the case in regions where one residential or business address is assigned a unique zip code for ease in mail delivery. ZCTA-level data is commonly used in the social sciences to measure neighborhood-level structural factors including socioeconomic status, unemployment rate, percent of vacant housing, segregation, and healthcare utilization.<sup>144-147</sup> ZCTA-level percentage of male-male unmarried partners has also been used to model "gay neighborhood" associations with health.<sup>148</sup>

Neighborhood-level structural stigma. To generate a treatment variable capturing exposure to neighborhood-level structural stigma, the percent of 'yes' votes for Prop 8 in a ZCTA (i.e., proportion of neighborhood voting in favor of a same-sex marriage ban) was generated, and then dichotomized into a binary category: 0 = less than 50% of ZCTA voted in favor of Prop 8 ('untreated'); 1 = 50% or greater of ZCTA voted in favor of Prop 8 ('treated'). Public precinct-level voting data for Prop 8 were downloaded from the California Secretary of State website. Precinct-level data were then aggregated to the ZCTA level using crosswalk conversion files that are available through the Statewide Database, the redistricting database for the state of California.

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<sup>&</sup>lt;sup>a</sup> See: https://www.sos.ca.gov/elections/prior-elections/statewide-election-results/

b See: http://statewidedatabase.org/fag.php?category=Conversion%20Files

The Prop 8 vote occurred in 2008; however, treatment status was applied to participants across all survey years (2004 – 2012). Gallup poll historical trend data demonstrates that national public approval of homosexuality and same-sex marriage has steadily increased over time. For instance, in response to the question, Do you think gay or lesbian relations between consenting adults should or should not be legal? 52% of respondents stated that this should be legal in May 2004, 55% in May 2008, and 63% in May 2012. Given this upward linear trajectory, with 2008 being an approximate, but slightly underestimated, average of public opinion in 2004 and 2012, I assumed for these analyses that the treatment categorization of neighborhood-level structural stigma in 2008 was an adequate approximation of average stigma across survey years.

Final Analytic Sample

ZCTAs were missing for 115 participants who were dropped from the analysis; thus the final analytic sample included 7,421 participants. A total of 442 unique ZCTAs were represented in the analysis. The weighted percent of 'yes' votes on Prop 8 across ZCTAs represented in the dataset was 52.6% (Std Err: 0.26). For comparison, the true proportion of 'yes' votes for Prop 8 for the entire state of California was 52.3%, thus demonstrating the representativeness of this survey sample to the state of California. The 'untreated' group included 3,265 participants (Wt % = 37.8%), and the 'treated' group included 4,156 participants (Wt % = 62.2%).

### Individual Variables

Sexual orientation was captured in Cal-QOL by asking participants if they self-identified as lesbian/gay/homosexual, bisexual or heterosexual. As well, participants were asked about the sex of their sexual partners since their 18<sup>th</sup> birthday. Participants who self-identified as

<sup>c</sup> See: https://news.gallup.com/poll/1651/gay-lesbian-rights.aspx

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lesbian/gay/homosexual or bisexual or who had had same-sex sexual partners since their 18<sup>th</sup> birthday were classified as a sexual minority.

Gender was recorded as male or female by the interviewer, If the interviewer was unsure, participants were asked to identify their gender as "male", "female", or "neither – transgender". No participants in the sample were identified as transgender.

Age was measured in years in the Cal-QOL. Age is categorized into the following categories: 18-29, 30-39, 40-49, 50-59, and 60+. The age range in the Cal-QOL is 18 to 77.

Race/ethnicity was classified in the Cal-QOL through a number of questions relating to racial/ethnic origin, including whether or not the participant identified as Hispanic/Latino. From this, respondents' race/ethnicity was categorized into one of five groups: Non-Hispanic White, Hispanic/Latino, Black/African-American, Asian/Pacific Islander, and American Indian/Alaska Native.

*Marital/Cohabitation status* was assessed by the following question: "Are you now married, living with a partner in a marriage-like relationship, widowed, divorced, separated, or never married?" Participants were classified as married/cohabiting or other.

Income and Federal Poverty Level were determined by asking participants their "household's annual income from all sources before taxes." Income was classified into the following four categories: less than \$20,000, \$20,000-\$59,999, \$60,000-\$99,999 and \$100,000 or greater. These cutoff points provided an even distribution across categories with a median cut point similar to the median household income in California in 2008 (~\$60,000). To provide a measure of income adjusted for inflation across survey years, a binary cutoff for Federal Poverty Level (FPL) was derived from income and household size. This variable was dichotomized as: 0 = income greater than 200% of the FPL, and 1= income less than 200% of the FPL.

### Time Variable

Four time points  $(t_1 - t_4)$  were utilized to examine the immediate and long-term impact of the Prop 8 vote on psychological distress among sexual minorities. The time points were based on the Cal-QOL collection waves with the second and third time points separated by the Prop 8 vote:  $t_1$ : October 2004 – February 2005 (Cal-QOL I respondents; n=2,245);  $t_2$ : August 2008 – November 3<sup>rd</sup> 2008 (Cal-QOL II respondents; n=2,135),  $t_3$ : November 4<sup>th</sup> 2008 – January 2009 (Cal-QOL II respondents; n=636), and  $t_4$ : January 2012 – February 2013 (Cal-QOL III respondents; n=2,405).

### Statistical Method

## Analytic Approach

A difference-in-differences (DID) approach was utilized to measure the treatment effect of neighborhood-level structural stigma on psychological distress among sexual minorities in the wake of the Prop 8 vote. DID is a quasi-experimental design that uses observational, longitudinal, or cross-sectional panel data from treated and untreated groups to estimate the causal effect of a treatment or intervention. <sup>150</sup> In econometrics and health policy, DID is common in measuring the effects of policy changes on populations in which a true control group is not available. <sup>151</sup> In social sciences, it is sometimes referred to as an "untreated control group design with pretest and posttest samples." <sup>152</sup> In its most common form, DID is utilized to measure prepost differences in the outcome of interest utilizing two groups; one group that received some policy change or intervention and another group that did not receive some policy change or intervention. The current analysis expands upon the commonest approach to allocate treatment group status based on the intensity of exposure to a shock (i.e., the 'intensity' of local-level stigmatization as measured by percent of one's neighbors voting to ban same-sex marriage). This

unique application of DID has been successfully utilized in a prior study measuring geographic variation in the intensity of an earthquake on maternal birth outcomes. <sup>153</sup>

DID is commonly implemented as an interaction term between time and treatment group dummy variables in a regression model. <sup>154</sup> In its most general iteration, the outcome  $Y_i$  is modeled by the following equation:

$$Y_i = \beta_1 * treated_i + \beta_2 * time + \beta_3(treated_i * time) + X_i$$

where the  $\beta_3$  coefficient for the interaction term captures the pre-post difference in average outcome in the *treated* group minus the pre-post difference in the average outcome in the *untreated* group. The  $\beta_1$  coefficient measures the average permanent differences between the treated and untreated groups. The  $\beta_2$  coefficient is a time trend common to the treated and untreated groups.  $X_i$  captures a set of individual-level covariates.

To capture how sexual minorities might be affected by the Prop 8 vote as compared to heterosexuals, I employed a triple differences or difference-in-difference-in-differences (DDD) methodology. With this methodology, a sexual minority subgroup is defined (i.e., those primarily affected by Prop 8) versus a comparison subgroup of similarly located heterosexuals who were expected to be less affected by the Prop 8 vote, and a three-level interaction term is added to the standard DID regression equation. Thus, the resulting DDD model is as follows:

$$\begin{aligned} Y_i &= \beta_1 * treated_i + \beta_2 * subgroup_i + \beta_3 * (treated_i * subgroup_i) + \beta_4 * time \\ &+ \beta_5 * (time * treated_i) + \beta_6 * (time * subgroup_i) \\ &+ \beta_7 * (time * treated_i * subgroup_i) + X_i, \end{aligned}$$

where the  $\beta_7$  coefficient is the expected "treatment effect coefficient" in the subgroup of interest. Similar to multilevel modeling, DID estimation is designed to control for systematic differences across treatment groups that might emerge from neighborhood composition,

socioeconomic status, and other unknown structural-level confounders.<sup>153</sup> Models also adjusted for individual-level confounders (gender, age, race/ethnicity, marital status, and federal poverty level) that are associated with sexual orientation and psychological distress, and were weighted to account for the complex survey design of the Cal-QOL (see *Weighting* section below for a detailed description of weighting strategies).

The most critical assumption of the DID methodology is the parallel trend or common trend assumption. This assumption implies that if the treatment group had not been treated (i.e., if individuals living in stigmatizing neighborhoods had never experienced the Prop 8 vote), then the trend of psychological distress as compared to psychological distress among individuals living in non-stigmatizing neighborhoods would have been constant over time. **Figure 2.3** presents a line graph of raw data from the Cal-QOL showing trends in psychological distress across time by treatment group, demonstrating that the data adheres to the parallel trend assumption (i.e. in the time period before the Prop 8 vote, the trends in outcome by treatment group and subgroup status are parallel).

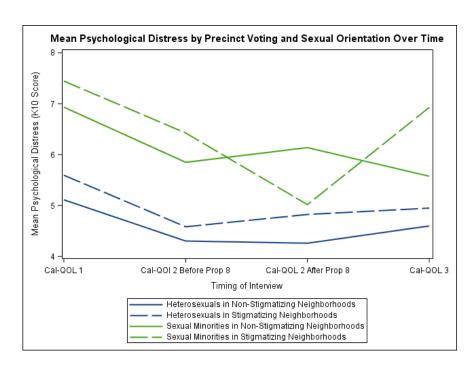


Figure 2.3. Unweighted trend of psychological distress across time by treatment group, California Quality of Life Survey

# Weighting

Two approaches to weighting were conducted to ensure that the analyses adequately accounted for the oversampling survey design of the Cal-QOL and to ensure that the 'exchangeability' assumption of causal inference was met (i.e., treated and untreated groups are equal in respect to characteristics except for the outcome of interest). If individuals living in neighborhoods that voted in favor of Prop 8 ('treated') and individuals living in neighborhoods that voted against Prop 8 ('untreated') are not comparable with respect to individual-level characteristics, then the estimation of differences in psychological distress will be biased. Thus, I first used the original Cal-QOL survey weights and, next, I incorporated the original survey weights with a propensity score method approach. Recent guidelines have suggested that, when applying propensity score methods to analyses of complex survey data, analyses should also incorporate the original survey weights or risk generating results that do not generalize to the

survey target population.<sup>156</sup> DuGoff et al. provide a decision tree to assist researchers in deciding what type of weighting is necessary based on the estimand of interest and the desired generalizability. Because the goal of the current analysis was to make a population-level inference (i.e., population average treatment effect), and to generalize to the survey target population (i.e., the state of California), I incorporated the original survey weights in the final analyses along with using propensity score methods.

To create a weight that incorporated propensity score methods and original survey weighting, I conducted a four-step process based on guidelines outlined by DuGoff et al. <sup>156</sup> First, I generated propensity scores by regressing individual-level covariates *and* the original Cal-QOL survey weight on treatment status. Second, propensity score balance was assessed by visual inspection of a histogram demonstrating propensity score distribution by treatment status (see **Figure 2.4**). Once a propensity score is calculated, one must determine that the range of propensity scores across treatment groups overlaps and has similar 'balance' in treated and untreated groups. <sup>157</sup> As demonstrated by Figure 2.4, the overlap and balance of propensity scores in the treated and untreated groups were satisfactory.

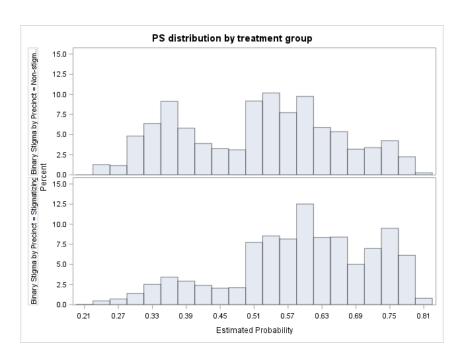


Figure 2.4. Propensity score distribution by treatment group demonstrating propensity score overlap and balance between treated and untreated groups

In a third step I calculated inverse probability of treatment weights (IPTW) from the propensity scores. IPTW produces unbiased estimates of the average treatment effect by generating a synthetic population composed of individuals in the original population weighted by the inverse of their probability of treatment; thus, the distribution of their baseline covariates is independent of treatment assignment. With IPTW, the treated group is weighted by  $\frac{1}{PS}$ , with PS being the propensity score, and the untreated group is weighted by  $\frac{1}{1-PS}$ . Finally, in a fourth step, I multiplied the original Cal-QOL survey weights by the IPTW to create final 'IPTW+original survey weights' that were utilized in final analyses, as described below. *Analyses* 

*Primary analyses*. All analyses were conducted in SAS 9.4. Weighted univariate analyses were conducted to characterize sample demographic characteristics. Then, bivariate analyses were employed to assess demographic characteristics between treatment groups under the

original weighting and propensity score weighting. Linear regression difference-in-difference-in-differences models, weighted by original survey weights and 'IPTW+original survey weight,' were then conducted to assess continuous K-10 score outcomes. Finally, logistic regression difference-in-difference-in-differences models with both weighting schemes were used to assess a binary outcome of 'meeting criteria for any K-10 distress' utilizing the following cutoff: 0, no distress: score 0-9 on the K-10; 1, any distress: score 10+ on the K-10.  $^{139,140}$  Models also adjusted for gender, age, race/ethnicity, marital status, and federal poverty level. Results were evaluated at  $p \le 0.05$ . All confidence intervals (CI) are 95% CI. This study was approved by the Institutional Review Board at the University of California, Los Angeles.

Supplemental Analyses. Several additional analyses were conducted to triangulate results from the primary analyses and elicit further analytic insights. First, I conducted falsification tests to assess the association between treatment status and health outcomes that would not be expected to be associated with the Prop 8 vote, including the presence of (i) asthma and (ii) heart disease. If treatment status was associated with either or both of these health outcomes, it would suggest that uncontrolled confounders influencing poor health outcomes were driving the obtained results.

Second, two robustness checks were conducted to test the sensitivity of the findings: (i) A three-level treatment status variable was created to test whether an even greater intensity of the local neighborhood environment would influence results. The percent of a neighborhood voting 'yes' on Prop 8 was trichotomized into tertiles representing the lower third (<43.87% voted 'yes' on Prop 8), middle third (43.87% - 61.02% voted 'yes' on Prop 8), and highest third (>61.02% voted 'yes' on Prop 8) levels of structural stigma and this three-level variable was included as the treatment variable in the DDD models. (ii) A treatment variable was created removing

participants living in neighborhoods with the least stigmatizing ( $5^{th}$  percentile;  $\leq 14.14\%$  voted 'yes' on Prop 8) and most stigmatizing ( $95^{th}$  percentile;  $\geq 75\%$  voted 'yes' on Prop 8) constituents (N = 6,689). This second robustness check sought to determine whether individuals living in highly-anomalous environments were driving results.

Third, because the results from these analyses (explained below) were contradictory to the primary hypothesis, I conducted post-hoc explanatory analyses to investigate why the findings were not in the predicted direction based on the prevailing structural stigma literature. Because of the robust literature documenting the protective role of social support in mitigating psychological distress among sexual minorities, 160-165 I conducted a set of analyses to identify whether social support differed by treatment status, time point, and sexual minority status, and might plausibly explain the discordant findings. For these analyses, I drew on a measure of social support included in the Cal-QOL. Six items in the Cal-QOL, drawn from the Medical Outcomes Study Social Support Survey, 166 measured social support in the past 4 weeks on a scale of 1 (all of the time) to 5 (none of the time), with items asking about social support in the domains of help ("how often was there someone who could help with daily chores if you were sick?"), love ("how often was there someone to love and make you feel wanted?"), conversation ("how often was there someone to confide in or talk to about yourself or your problems?"), time ("how often was there someone to have a good time with?"), information ("how often was there someone to give you information to help you understand a situation?"), and money ("how often was there someone to give you money if you needed it?"). Data were missing on at least one of the social support items for 7.1% of the sample. Individual social support items were imputed with single imputation methods; then, items were reverse-coded, summed, and rescaled so that the social

support scale score ranged from 0 to 24, with 24 being the highest possible level of social support. The weighted mean score in the total sample was 18.81 (Range: 0 - 24; Std Err: 0.08).

## Results

Sample Characteristics

I analyzed data from 7,421 individuals from the California Quality of Life Survey interviewed in 2004-2005, 2008-2009, and 2012-2013. Univariate weighted averages were utilized to characterize sample demographics. Sexual minorities comprised 3.08% (SE: 0.12) of the sample as compared to 96.92% (SE: 0.12) heterosexuals. A slight majority of participants were female 52.02% (SE: 0.76) as compared to male 47.97% (SE: 0.76). The mean age of the sample was 42.5 years (SE: 0.22). Participants were predominately Non-Hispanic White (49.44%; SE: 0.76), followed by Hispanic/Latino (32.6%; SE: 0.74), Asian/Pacific Islander (10.23%; SE: 0.53), Black/African-American (6.83%; SE: 0.32), and American Indian/Alaska Native (0.86%; SE: 0.13). About one-fifth (21.56% [SE: 0.64]) of the sample earned less than \$20,000 per year; 31.1% (SE: 0.72) of participants earned less than 200% of the Federal Poverty Level. Almost two-thirds of the sample were married or cohabiting (62.52%; SE: 0.74). Allocation of treatment status by neighborhood voting characteristics demonstrated that 37.8% (SE: 0.74) of the sample resided in a neighborhood that voted against Prop 8, ('untreated') and 62.2% (SE: 0.74) resided in a neighborhood that voted in favor of Prop 8 ('treated').

Table 2.1. Sample demographic characteristics by neighborhood voting status, California Quality of Life Survey 2004-2012 (N=7,421)

	Original su	Original survey weights		IPTW + Orig	IPTW + Original survey weights	
	Neighborhood voted against same-sex marriage ban (n=3,265)	Neighborhood voted in favor of same-sex marriage ban (n=4,156)		Neighborhood voted against same-sex marriage ban (n=3,265)	Neighborhood voted in favor of same-sex marriage ban (n=4,156)	
Demographic Characteristics	Wt % (SE)	Wt % (SE)	p-value <sup>a</sup>	Wt % (SE)	n (Wt %)	p-value <sup>a</sup>
Sexual Orientation Heterosexual Sexual Minority	95.5 (0.2) 4.5 (0.2)	97.8 (0.1) 2.2 (0.1)	<.001	97.1 (0.2) 2.9 (0.2)	96.5 (0.2) 3.5 (0.2)	0.02
<b>Gender</b> Female Male	50.4 (1.2) 49.6 (1.2)	53.0 (1.0) 47.0 (1.0)	ns	52.2 (1.3) 47.8 (1.3)	51.7 (1.0) 48.3 (1.0)	ns
Age (M, SE)	43.2 (0.4)	42.2 (0.3)	0.02	42.5 (0.4)	42.6 (0.3)	Su
Race/ethnicity Non-Hispanic White Non-Hispanic Black Hispanic/Latino Asian/Pacific Islander American Indian/Alaska Native	56.3 (1.2) 6.7 (0.5) 22.7 (1.1) 13.5 (1.0) 0.8 (0.2)	45.3 (1.0) 6.9 (0.4) 38.7 (1.0) 8.3 (0.6) 0.9 (0.2)	<.001	48.5 (1.4) 7.4 (0.6) 32.4 (1.4) 10.6 (0.8) 1.0 (0.3)	50.0 (0.9) 6.6 (0.4) 32.6 (0.9) 10.0 (0.7) 0.8 (0.2)	ns
Married/cohabiting Yes No	61.4 (1.2) 38.6 (1.2)	63.2 (0.9) 36.8 (0.9)	su	61.9 (1.3) 38.1 (1.3)	62.6 (0.9) 37.4 (0.9)	ns
Income <\$20,000 \$20,000-\$59,999 \$60,000-\$99,999 \$100,000+	18.1 (1.0) 29.9 (1.1) 21.5 (1.0) 30.5 (1.1)	23.6 (0.8) 36.6 (0.9) 20.3 (0.8) 19.5 (0.8)	<.001	22.6 (1.2) 32.7 (1.3) 20.5 (1.0) 24.2 (1.0)	21.3 (0.8) 34.2 (0.9) 21.0 (0.8) 23.5 (0.9)	ns

10.00 T		100 /			SU
>700 % OI reaeral roverty Level		100.			
Yes	23.8 (1.1)	35.6 (0.9)	29.2 (1.3)	32.1 (0.9)	
No	76.2 (1.1)	64.4 (0.9)	70.8 (1.3)	(6.0)(6.2)	
Note. Wt = Weighted percentages.					
<sup>a</sup> P-value evaluated by t-test or $\chi^2$ test as	аррі	opriate. $ns = \text{not significant}$ ; $p \ge 0.05$ .			

Table 2.1 shows results from bivariate analyses demonstrating differences between groups based on treatment status (i.e., neighborhood in which an individual resides). In the overall weighted sample (original weights), there are several significant differences by treatment group status, including that as compared to individuals living in neighborhoods that voted in favor of Prop 8, individuals living in neighborhoods that voted against Prop 8 were slightly older, more likely to be a sexual minority, more likely to be non-Hispanic White and less likely to be Hispanic/Latino, and more likely to earn a higher annual income. After propensity score weighting (IPTW+ original weights), individuals living in 'treated' and 'untreated' groups were balanced with respect to covariates. As demonstrated by Table 2.1, there were no significant differences between treatment groups with respect to any demographic characteristics except for sexual orientation where propensity weighting served to reduce the absolute difference in the weighted proportion of sexual minorities between the groups (2.9% vs. 3.5%).

Primary Analyses

**Table 2.2** presents results from the primary difference-in-difference-in-differences analyses for the outcome of psychological distress (K-10 score) as measured: (i) on a continuous scale and (ii) with a binary cutoff. The partial results of four models are shown; one for each outcome (linear or logistic) with each weighting scheme (original or incorporating propensity score weights). The parameter estimate of interest for each of the models is the interaction term between time period, treatment group status, and the subgroup of interest (i.e., sexual minorities). Model 1 predicts mean psychological distress across time by treatment status and subgroup of interest and is weighted with original Cal-QOL survey weights. Results demonstrated that sexual minorities living in neighborhoods where a majority of constituents voted in favor of Prop 8 (a same-sex marriage ban; i.e., 'treated') had a statistically significant mean psychological distress

reduction of 4.06 on the K-10 scale (95% CI: -7.38; -0.73) in the time period immediately after the Prop 8 vote. In Model 3, the integration of propensity score methods demonstrated little change in this result, with only a slight attenuation to a statistically significant mean psychological distress reduction of 3.71 (95% CI: -6.85; -0.56). Logistic regression models (Models 2 and 4) demonstrated parallel results to the linear regression models, highlighting that not only was there a significant reduction in mean psychological distress in the period after the Prop 8 vote among 'treated' sexual minorities, but also that this reduction in psychological distress was associated with significantly reduced odds of meeting the threshold for mild, moderate, or severe psychological distress. Model 4 also demonstrated a significant increase in psychological distress in the follow-up time period among sexual minorities in the 'treated' group (aOR = 1.22; 95% CI: 1.02-1.45).

Table 2.2. Partial results from difference-in-difference-in-differences analyses assessing the effects of neighborhood voting status on psychological distress: California Quality of Life Survey, 2004-2012 (N = 7421)

			Original Survey Weighting	ey Weigh	ting			1	IPTW+Original Survey Weighting	Survey	Weighti	g <sub>1</sub>
	X	Model 1: F Conti	lel 1: K-10 Score, Continuous <sup>a</sup>	Mod	el 2: K-10 Disor Cutoff, Binary <sup>b</sup>	Model 2: K-10 Disorder Cutoff, Binary <sup>b</sup>	Mo	del 3: K Contir	Model 3: K-10 Score, Continuous <sup>a</sup>	Mod	el 4: K-10 Dison Cutoff, Binary <sup>b</sup>	Model 4: K-10 Disorder Cutoff, Binary <sup>b</sup>
	Coef.	SE	95% CI	aOR	SE	95% CI	Coef.	SE	95% CI	aOR	SE	95% CI
Heterosexual [ $Subgroup = 0$ ]												
Sexual Minority [ $Subgroup = 1$ ]	2.33	0.58	(1.20; 3.47)	1.32	0.07	(1.16; 1.50)	2.37	0.62	(1.16; 3.57)	1.35	0.07	(1.19; 1.54)
Less than 50% of Neighborhood voted in favor of Same-Sex Marriage (SSM) Ban $[V = 0]$												
50% or More of Neighborhood voted in favor of Same-Sex Marriage (SSM) Ban $[V = 1]$	0.24	0.31	(-0.36; 0.86)	66.0	90.0	(0.88; 1.13)	0.17	0.33	(-0.48; 0.82)	0.98	90.0	(0.86; 1.11)
Baseline (2004-2005) [ <i>t</i> =0]												
3 Months Preceding Same-Sex Marriage Ban [f=1]	-0.42	0.34	(-1.10; 0.22)	1.14	0.09	(0.95; 1.37)	-0.53	0.37	(-1.24; 0.18)	1.19	0.10	(0.98; 1.45)
3 Months Following Same-Sex Marriage Ban [ <i>t</i> =2]	-0.94	0.44	(-1.80; -0.07)	0.77	0.15	(0.57; 1.03)	-1.09	0.46	(-1.99; -0.19)	92.0	0.15	(0.56; 1.02)
Follow-Up (2012-2013) [ <i>t</i> =3]	-0.34	0.34	(-1.00; 0.32)	96.0	60.0	(0.80; 1.14)	-0.77	0.44	(-1.06; 0.46)	96.0	60.0	(0.80; 1.15)
$Subgroup_I(Sexual Minority) \times V_I(Pro-SSM Ban) \times t_0(Baseline)$												
$Subgroup_I(Sexual Minority) \times V_I(Pro-SSM Ban) \times t_I(Before SSM Ban)$	-0.54	1.48	(-3.44; 2.37)	1.16	0.09	(0.96; 1.39)	-0.84	1.55	(-3.87; 2.20)	1.15	0.10	(0.95; 1.40)
$Subgroup_I(Sexual Minority) \times V_I(Pro-SSM Ban) \times t_2(After SSM Ban)$	-4.06	1.69	(-7.38; -0.73)	0.70	0.15	(0.52; 0.95)	-3.71	1.61	(-6.85; -0.56)	69.0	0.15	(0.52; 0.94)
Subgroup <sub>1</sub> (Sexual Minority) x $V_1$ (Pro-SSM Ban) x $t_3$ (Follow-up)	1.20	1.43	(-1.60; 3.99)	1.19	0.09	(0.99; 1.41)	1.49	1.35	(-1.14; 4.13)	1.22	0.09	(1.02; 1.45)

Intercept	6.07	0.36	(5.36; 6.78)	0.28	0.10	(0.23; 0.34)	0.09	0.39	(5.23; 6.76)	0.27	0.10	(0.22; 0.33)	
Note M = 7.421 Models weighted and adjusted	dineted for	T Age G	Gender Married/co	habiting	Ctatue D	Pace and Federal	Doverty	I eyel					

Note. N = 7,421. Models weighted and adjusted for Age, Gender, Married/cohabiting Status, Race, and Federal Poverty Level. IPTW = Inverse Probability of Treatment Weighting; Coef. = Coefficient; aOR = adjusted odds ratio, 95% CI = 95% Confidence Intervals Statistical significance **bolded** at p ≤0.05.

<sup>a</sup>K-10 Score range: 0 - 40 <sup>b</sup>K-10 Disorder Scaling: Score 0-9; no distress (0), Score 10 or higher; any distress (1)

Supplementary Analyses – Robustness Checks

Falsification tests were run by regressing the difference-in-difference-in-differences models from the primary analyses on the outcomes of the presence of (i) asthma and (ii) heart disease. Results revealed, as expected, that treatment status was not associated with these health outcomes that theoretically should not be affected by neighborhood voting behavior.

Sensitivity analyses included running the difference-in-difference-in-differences models from the primary analyses with (i) a three-level treatment variable (i.e., lowest, middle, highest percent of constituents voting 'yes' on Prop 8) and (ii) with the two-level treatment variable excluding individuals living in neighborhoods that represented the 5<sup>th</sup> and 95<sup>th</sup> percentile of constituents voting 'yes' on Prop 8 (i.e., included neighborhoods with 14.14% - 75% 'yes' votes). Linear regression DDD models utilizing the three-level treatment variable with original Cal-QOL survey weights and continuous K-10 outcome demonstrated that the influence of neighborhood voting behavior on psychological distress was even more severe when voting behavior was trichotomized into tertiles. Indeed, the parameter estimate from the desired interaction term,

 $\beta_7(Subgroup_{Sexual\ Minority}*Time_{After\ Prop\ 8}*Treatment_{Highest\ Stigma})$  demonstrated a statistically significant mean psychological distress reduction of 5.70 (95% CI: -9.72; -1.68). This highlights that, as compared to sexual minorities living in neighborhoods with the least support for Prop 8, sexual minorities in neighborhoods with the most support for Prop 8 saw a significant psychological distress reduction in the period after the vote. The second robustness check excluded participants living in neighborhoods with voting behavior representing the 5<sup>th</sup> and 95<sup>th</sup> percentiles of voting behavior (excluded n = 732). Results showed little difference from the primary analyses, with the parameter estimate of interest showing a

statistically significant mean psychological distress reduction of 3.53 (95% CI: -7.02; -0.04), confirming that these results are not being driven by individuals living in anomalous neighborhood environments.

Additional Analyses – Social Support

A series of post-hoc exploratory analyses were conducted through an iterative process to investigate the role of social support in influencing the aforementioned unexpected results. First, to obtain a visual depiction of trends in social support across time, I graphed unweighted mean social support by treatment group, subgroup status, and time. From this graph (see **Figure 2.5**) it was evident that mean social support indeed differed across time and that, in particular, there was an increasing trend in mean social support among sexual minorities living in 'treated' (i.e., stigmatizing) neighborhoods in the time period immediately after the Prop 8 vote. I then conducted a series of bivariate and multivariate analyses to investigate the significance of this trend.

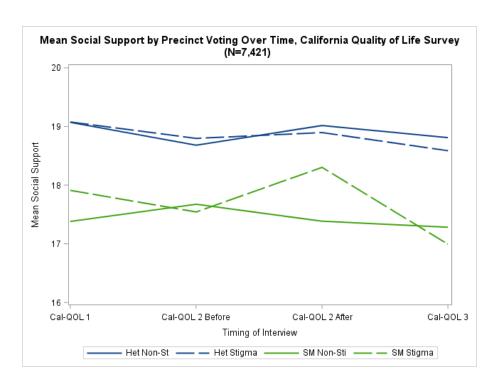


Figure 2.5. Unweighted mean social support across time by treatment group ('stigma' vs. 'non-stigma') and subgroup ('heterosexual' vs. 'SM [sexual minority]') status

Next, I investigated bivariate differences in social support by sexual minority status and treatment group status. In the total weighted sample, sexual minorities reported significantly lower mean social support (M = 17.66, Std Err: 0.23) as compared to heterosexuals (M = 18.85; Std Err: 0.08); p < .0001). There was no evidence of mean difference in social support by neighborhood voting behavior (treatment group). As well, when restricting bivariate analyses to sexual minorities, there was no significant main effect difference in mean social support by treatment group status.

I then ran a series of weighted linear regression models restricted by treatment group status at each time point to investigate whether the significant subgroup differences (i.e., sexual minority vs. heterosexual) in average social support remained across time. Models adjusted for age, race, and sex. When restricting these analyses to 'untreated' participants (i.e., residing in neighborhoods where less than 50% of constituents voted in favor of Prop 8), sexual minorities

demonstrated significantly lower average social support than heterosexuals at each of the four time points (p < .05). However, when restricting analyses to 'treated' participants (i.e., residing in neighborhoods where greater than or equal to 50% of constituents voted in favor of Prop 8), average social support among sexual minorities was significantly lower as compared to heterosexuals at time points 1 (Cal-QOL Wave 1) and 4 (Cal-QOL Wave 4), but the adjusted mean social support difference between sexual minorities and heterosexuals was not significant at time points 2 and 3 (Cal-QOL Wave 2 Before Prop 8 and Cal-QOL Wave 2 After Prop 8), tentatively suggesting that mean social support may have differed somewhat across time and treatment group status.

To test the causal impact of the Prop 8 vote on social support among sexual minorities, I ran the weighted difference-in-difference-in-differences linear regression models from the primary analyses with social support as the outcome of interest. The obtained parameter estimate for the  $\beta_7$  coefficient of interest was 2.79 (95% CI: -0.22; 5.80), highlighting that while this did not reach significance at p  $\leq$  .05, the parameter estimate was positive, in the expected direction, and the 95% confidence interval demonstrated that the parameter trended towards a positive mean increase in social support among sexual minorities living in 'treated' neighborhoods in the time period immediately after the Prop 8 vote (p = 0.07). These results are not confirmation that social support was a reason for the unexpected results from the primary analyses (that sexual minorities living in treated neighborhoods reported *reduced* psychological distress after the Prop 8 vote); rather, these post-hoc analyses provide exploratory evidence that social support may indeed have been one factor influencing these findings. Only further analyses with additional measures of social support not available within the current study can help to triangulate these preliminary findings; indeed, future studies should seek to investigate the causal role of social support as a

buffer against psychological distress among sexual minorities in stigmatizing environments and in the face of discriminatory policies.

#### **Discussion**

Findings from this representative study of 7,421 adults in California demonstrated that a state-level vote to effectively ban same-sex marriage was associated with a significant *reduction* in psychological distress among sexual minorities living in neighborhoods in which a majority of constituents voted in favor of the proposition as compared to those living in neighborhoods in which a majority of constituents voted against the proposition. Several robustness checks were conducted to triangulate these unexpected findings. The unanticipated results from this study complicate understanding how structural stigma operates to produce psychological distress among sexual minorities. However, while these results do not fit neatly into the current understanding of the impact of structural stigma on sexual minority mental health, they do offer an opportunity to think critically about reactions to stigmatizing events and, especially, to consider how social support and other resilience factors can buffer the harmful effects of structural stigma.

Several studies have sought to investigate the association between anti-LGB amendments and mental health among sexual minorities. <sup>134,167-171</sup> These studies primarily drew on a minority stress framework to explain the increased levels of psychological distress experienced by sexual minorities following votes on anti-LGB amendments including same-sex marriage bans. However, most used convenience sampling strategies and cross-sectional data of individuals drawn often times from higher gay density environment. These designs limit causal inference from these studies. Nonetheless, in addition to documenting how anti-LGB policies can negatively impact mental health, many of these studies also reported how stigmatizing events

like a statewide vote to ban same-sex marriage can foster social support, LGB communitybuilding, and resilience among affected sexual minority populations. Indeed, Russell and Richards (2003) documented resilience factors among 316 sexual minorities following an amendment to Colorado state's constitution to deny legal recourse to LGB people who had experienced sexual orientation-related discrimination. <sup>171</sup> Five main resilience factors emerged from that research, including (1) movement perspective (e.g., "Increased LGB community's visibility through publicity and media coverage"), (2) confronting internalized homophobia (e.g., "Felt less shame as a LGB person"), (3) expression of affect (e.g., "Had opportunity to express anger"), (4) successful witnessing (e.g., "Heterosexual friends offered understanding and support"), and (5) LGB community (e.g., "Increased support from LGB community"). A similar study that sought to document reactions to California's Prop 8 – the same-sex marriage proposition investigated in the current study – among 354 sexual minorities recruited from LGBT venues, social media, and "No on Proposition 8" campaign events found that participants reported similarly moderate levels of positive (e.g., 'happy', 'pride') and negative (e.g., 'upset', 'angry') emotions. 134 Further, participants also reported experiencing statistically significantly greater support than conflict from their intimate partner, family, friends, coworkers, and the heterosexual community in relation to Prop 8. Results from the present study expanded on findings from these prior studies to demonstrate that psychological resilience might have been temporarily enhanced among sexual minorities in the wake of a stigmatizing same-sex marriage ban. This finding might be partially explained by some of the aforementioned resilience factors: indeed, sexual minorities living in neighborhoods where a majority of constituents voted in favor of a same-sex marriage ban might have positively benefited from these resilience factors – especially social support – more so than sexual minorities in less stigmatizing (i.e., more

supportive) neighborhoods where these resilience factors may have already been more common, or even the status quo. Post-hoc analyses in the current study showed that perceived social support among sexual minorities was significantly lower as compared to heterosexuals at each time point among those living in 'non-stigmatizing' neighborhoods, but not significantly different at time points 2 and 3 (immediately before and after Prop 8) among those living in 'stigmatizing' neighborhoods. Thus, for sexual minorities in 'stigmatizing' neighborhoods, their perception of levels of social support in the time period immediately before and after the Prop 8 vote was on par with the perceived levels of social support among heterosexuals. This uptick in perceived social support might partially explain the reduction in psychological distress among this group in the wake of the same-sex marriage ban, and it highlights how perceptions of social support in the sample differed by sexual minorities' local residential environments.

Prior research investigating how geography moderates perceived social support and access to resources among sexual minorities has primarily focused on the urban-rural divide: some studies have demonstrated that sexual minorities in rural locales are less connected to LGB communities, experience more discrimination, have significantly lower social engagement, and are less likely to be 'out' to family and social circles as compared to those in urban communities. However, other studies — mainly qualitative investigations and ethnographies — have demonstrated that sexual minorities in rural communities often construct their queer identities in concert with their identity as a "local," incorporating strong ties to their local, 'small town' communities in tandem with building 'gay community' support through online venues. 174-176 While the current study did not investigate an urban-rural divide, but rather an electoral divide, the reduction in psychological distress among sexual minorities in stigmatizing neighborhoods after the Prop 8 vote elucidates how stigmatizing events can differentially impact

individual-level mental health among sexual minorities depending on the context of where they live. Future investigations – both qualitative and quantitative – should attempt to unpack how sexual minorities' local environments moderate reactions to events that are objectively stigmatizing to increase our understanding of how stigma might operate differently than we presume and how resilience can be drawn on by sexual minorities to buffer the harmful effects of discriminatory policies and events.

### **Limitations and Conclusions**

This study is subject to a number of potential limitations. First, I operationalized the proportion of votes for a same-sex marriage ban in a neighborhood as an objective measure of structural stigma in that neighborhood. This strategy rests on the assumption that there were no other paths of influence, other than stigmatization, whereby exposure to the Prop 8 vote had a detrimental effect on psychological distress among sexual minorities. Further, this strategy also assumed that sexual minorities in the neighborhood were aware of their neighbors' views of their right to marry beyond just the state-level Prop 8 vote. While studies have highlighted that Prop 8 was highly advertised across mediums, including news and local advertisements, posters, and vard signs. 84,134 it is plausible that some sexual minorities were unaware of the support of Prop 8 within their own neighborhood. A second primary concern was that the Prop 8 vote occurred on the same day as the 2008 Presidential Election of Barack Obama. Thus, while Prop 8 was a salient, stigmatizing event in California for sexual minorities, the effect of the vote cannot be separated from the effects that the Presidential Election might have had on sexual minority psychological distress. Because sexual minorities evidence greater political liberalism than heterosexuals. 177 the election of Barack Obama, a Democrat, might have had a differential impact on psychological distress among sexual minorities as compared to heterosexuals. A third

limitation to this study is that the political engagement of participants was unknown: if, for example, sexual minorities in less stigmatizing environments were significantly more likely to engage in political activism, this might have confounded results. Last, the main outcome (K-10 psychological distress) was based on self-report, and might have been underreported given the sensitive nature of the questions and that the measure is collected via telephone rather than with a computer-based questionnaire.<sup>178</sup>

Despite these limitations, the study also demonstrated numerous strengths. A primary strength of this study was that it met the gold-standard methodology of employing a quasi-experimental design. Indeed, by using cross-sectional panel data across time and utilizing a causal difference-in-difference analytic strategy with propensity weighting, this study improves our understanding of temporal changes in psychological distress before and after a stigmatizing event. As well, the study participants were drawn from the Cal-QOL, a methodologically-rigorous statewide survey that is representative of the population of California. The only known prior study investigating the impact of Prop 8 on psychological distress among sexual minorities employed convenience sampling with a small sample, likely leading to selection biases. 134

Findings from this study showed that sexual minorities living in objectively stigmatizing neighborhoods – as indexed by the proportion of neighborhood constituents voting in favor of a proposition that effectively banned same-sex marriage – reported reduced psychological distress in the two-month period after the vote. This project adds to – and complicates – our understanding of how structural stigma operates to produce psychological distress among sexual minorities at a neighborhood-level; as well, these findings emphasize the need for additional research on resilience factors and protective influences, especially social support, in the wake of stigmatizing events.

# Appendix 2.A.

# **Kessler 10 Psychological Distress Scale**

The following questions ask about how you have been feeling during the **past 30 days**. For each question, please circle the number that best describes how often you had this feeling.

Q1. During that month, how often did you feel	All of the time	Most of the time	Some of the time	A little of the time	None of the time
a tired out for no good reason?	1	2	3	4	5
<b>b.</b> nervous?	1	2	3	4	5
<b>c.</b> so nervous that nothing could calm you down?	1	2	3	4	5
dhopeless?	1	2	3	4	5
erestless or fidgety?	1	2	3	4	5
fso restless that you could not sit still?	1	2	3	4	5
gdepressed?	1	2	3	4	5
hso depressed that nothing could cheer you up?	1	2	3	4	5
ithat everything was an effort?	1	2	3	4	5
jworthless?	1	2	3	4	5

# Chapter 3

Sexual orientation differences in gun ownership and beliefs about gun safety policy: Findings from the General Social Survey, 2010 - 2016

## Background

Gun violence in the United States is a major public health concern; gun-related injuries are estimated to cost \$2.8 billion annually in healthcare costs alone. 179 More than half of all violent deaths in 2012 in the U.S. involved the use of guns (69.6% of homicides; 50.9% of suicides). 180 In addition to mortality, in the first decade of the 21<sup>st</sup> century, the rate of non-fatal gun-related injury in the United States was twice the rate of fatal gun-related injury and accounted for 921,613 non-fatal injuries. 181 Gun ownership is also a crucial predictor of gun-related violence including suicide. 182 Access to a gun is a robust predictor of suicide beyond individual-level psychopathology. 34,183 In general, men are more likely to own guns than women, 42,182,184 which partially explains why men are more likely to die by suicide than women. 185 Whether differences in gun exposure are also shaped by sexual orientation is largely unknown, but may represent unrecognized differences in risk for suicide, a significant knowledge gap given sexual minorities' (e.g., lesbians, gay men, and bisexuals) disproportionate suicide risk. Blosnich et al. conducted the only known study of sexual orientation and gun ownership and found that samesex partnered individuals were significantly less likely to report having a gun in the household compared to opposite-sex partnered individuals, an association that was modified by military service history. 186 But this study was unable to test gender differences in sexual minority gun ownership, a critical limitation given differences in gender-based behavioral patterns among sexual minorities. 7,187,188

A related factor to gun ownership is personal views about gun safety laws. In general, sexual minorities voice more liberal political views as compared to heterosexuals;<sup>177</sup> thus, one might

expect that sexual minorities favor stricter gun control and gun safety laws. <sup>189</sup> After 49 lesbian, gay, bisexual and transgender (LGBT) individuals were killed in 2016 by an armed gunman at Pulse nightclub in Orlando, Florida, <sup>190</sup> activists, news media and policy centers argued that gun control must be viewed as an "LGBT policy issue" given the population's unique vulnerability to gun violence. <sup>189,191</sup> But, in actuality, no studies to date of which I am aware have investigated gun-specific policy views among sexual minority individuals.

Thus, I used information from multiple years of the General Social Survey to investigate sexual orientation-related patterns of gun ownership and beliefs about gun safety policy. I hypothesize that sexual minorities will be more endorsing of gun safety policy than their heterosexual peers. I also hypothesize that sexual minority men will have a lower prevalence of gun ownership as compared to heterosexual men. However, given evidence demonstrating the wide range of elevated behavioral health risks among sexual minority women (e.g., high rates of smoking, alcohol use, poor diet, lower rates of health care utilization <sup>192-194</sup>), I further hypothesize that sexual minority women will demonstrate a greater prevalence of gun ownership as compared to heterosexual women.

## **Public Health Significance**

Understanding the risk of gun ownership among sexual minorities might identify important and unrecognized pathways to reducing suicide risk in this vulnerable population. This study has the potential to identify specific public health intervention points related to sexual minority gun ownership. Interventions targeting safe gun ownership might be one important strategy in preventing suicide risk among sexual minorities. To that end, research has shown that means restriction – limiting a suicidal person's access to highly lethal means – is effective in suicide prevention. Means restriction includes training clinicians to ask about gun ownership and

counseling at-risk individuals and their families to store guns away from the home or making household guns inaccessible.<sup>34</sup> Thus, findings of this study might elucidate that mental health providers working with sexual minorities represent a key point of intervention in suicide prevention efforts. Inquiring about gun ownership and providing lethal means counseling to sexual minorities at risk for suicide would represent an innovative and important public health addition to mental health providers' suicide reduction toolkits.

### Methods

# Study Design and Population

The current study used cross-sectional data from multiple years of the General Social Survey (GSS). The GSS is a biennial survey that employs a multi-stage sampling design to obtain a sample of non-institutionalized adults (≥ 18 years old) living in the United States. For over four decades, the GSS has been a social litmus test: the survey includes items related to attitudes and beliefs on various topics including crime, politics, immigration, morality, national security priorities, and social mobility. Since 2002, the GSS has administered surveys to respondents using computer-assisted personal interviews. In 2008, the GSS added a question to gather respondents' self-identified sexual orientation. However, military veteran status, a strong modifier of gun ownership, <sup>195</sup> was not collected in that cycle; thus, the proposed project will analyze data from the GSS collected in 2010, 2012, 2014, and 2016. Response rates for the GSS ranged from 71.4% (in 2012) to 61.3% (in 2016).

From 2010-2016, three question ballots were implemented (A, B and C), and only certain questions were asked depending on ballot type (e.g., gun-related questions were not asked on ballot B). To determine a final analytic sample size, I conducted a set of pre-analytic data restriction measures. I restricted analyses to include only respondents who received all of the

following questionnaire items: gun safety policy beliefs, presence of guns in the household, sexual orientation, and an item asking about military veteran status (N=4,867). I dropped respondents with missing data for our key independent and dependent variables, including gun safety policy beliefs (n=74), presence of a gun in the household (n=155), and sexual orientation (n=139). Additionally, I removed respondents with missing data for military veteran status (n=7). Given the overlap in missing data, I dropped from consideration a total of 313 respondents. Thus, the final analytic sample for the proposed project includes 4,554 respondents who range in age from 18 to 89 years of age.

#### Measures

# Dependent Variables

I focused on three dependent variables: (1) beliefs about gun safety policy, (2) the presence of guns in the household, and (3) among those reporting household guns, whether the respondent was the gun owner. The following item measured gun safety beliefs: "Would you favor or oppose a law which would require a person to obtain a police permit before he or she could buy a gun?" Response options included "favor" or "oppose." The presence of gun(s) in the household was assessed by the question: "Do you happen to have in your home or garage any guns or revolvers?" This variable was coded trichotomously ("yes"/ "no"/ "refused"). If participants responded affirmatively, a follow-up question asked: "Do any of these guns personally belong to you?" This variable was coded dichotomously ("yes"/ "no").

# Primary Independent Variables

Respondents' gender was coded as male or female. The following item assessed sexual orientation: "Which of the following best describes you?" Response options included: "gay, lesbian, or homosexual", "bisexual", or "heterosexual or straight." Given the small number of

respondents who selected "gay, lesbian, or homosexual" or "bisexual," I combined these respondents into a single category denoting sexual minority status.

## Demographics

The GSS measures several other demographic characteristics including race/ethnicity, age, education level, household size, urbanicity, political views, and military veteran status. The GSS codes race/ethnicity in three categories (White, Black, or other). For analytic purposes, I recoded education level into two categories (high school or less versus more than high school); urbanicity into three categories (urban/suburban, small city/town, or rural); military veteran status into two categories (military veteran versus not); and political views into three categories (conservative, moderate, or liberal).

## Statistical Analysis

Statistical analyses were conducted in R taking into account the complex sampling design and weighting of the GSS using the "survey" package. Covariates with missing data included education (<1%); age (<1%); and political views (3%). Missing covariate data were imputed with Rubin's multiple imputation method<sup>196</sup> using the "mice" package in R.

Analyses were conducted with the full analytic sample as well as with gender-stratified (male only/female only) samples. Initially, bivariate analyses were used to determine unadjusted group differences between heterosexuals and sexual minorities for demographics characteristics and gun-related outcomes. I report the results of t-tests and Wald chi-square tests as appropriate. Then, multivariable logistic regression models were used to evaluate the adjusted association of sexual minority status with each of the three outcomes in the overall sample and in gender-stratified samples. Analyses controlled for several possible known confounders of gun safety beliefs and gun ownership including gender, age, race/ethnicity, education level, size of

household, urbanicity, military veteran status, and political views as well as survey year. <sup>195,197</sup> Results were evaluated at  $p \le 0.05$ . I report weighted percentages and standard errors (SE). All confidence intervals (CI) are 95% CI. This study was exempt from IRB review.

### Results

## Sample Characteristics

Approximately 3.6% (SE: 0.004) of the weighted GSS sample identified as a sexual minority (see **Table 3.1**). Among men, 3.5% (SE: 0.005) identified as a sexual minority (2.4% gay [SE: 0.004] and 1.2% bisexual [SE: 0.002]). Among women, 3.7% (SE: 0.004) identified as a sexual minority (1.1% lesbian [SE: 0.002] and 2.7% bisexual [0.004]). For both men and women, sexual minority status was significantly associated with younger age and more liberal political views. As well, sexual minority men reported higher education levels and smaller households when compared to heterosexual men. There was also a trend that sexual minority men were less likely to be military veterans as compared to heterosexual men ( $\mathbf{p} = 0.054$ ).

Table 3.1 Demographic characteristics among respondents in the 2010-2016 General Social Survey (GSS), by gender and sexual orientation

	To	Total		Wo	Women		M	Men	
	Sexual		ı	Sexual		I	Sexual		
	Minority	Heterosexual		Minority	Heterosexual		Minority	Heterosexual	
Characteristics	(n=195) Est (SE)	(n=4359) Fet (SE)	۵	(n=102) Fet (SF)	(n=2401) Fet (SE)	۵	(n=93) Fet (SF)	(n=1958) Fet (SF)	۵
Women %	55.6 (0.04)	54.2 (0.01)	0.71	1	1	4	(======================================	1	4
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Age, $\bar{x}$	38.8 (1.19)	46.4 (0.35)	<0.001**	37.1 (1.7)	46.7 (0.4)	<0.001**	41.0 (1.8)	46.0 (0.44)	$0.01^{*}$
Race/ethnicity, %			0.33			0.77			0.26
White	76.0 (0.04)	74.4 (0.02)		76.4 (0.05)	73.6 (0.02)		75.7 (0.06)	75.3 (0.02)	
Black	15.8 (0.03)	14.7 (0.01)		15.7 (0.04)	15.9 (0.02)		15.9 (0.04)	13.3 (0.01)	
Other	8.1 (0.02)	10.9 (0.01)		7.9 (0.03)	10.5 (0.02)		8.4 (0.03)	11.3 (0.01)	
Education Level, %			$0.04^*$			0.22			$0.03^{*}$
High school or less	31.4 (0.05)	41.5 (0.01)		33.5 (0.06)	40.7 (0.02)		28.9 (0.06)	42.4 (0.02)	
More than high school	68.6 (0.05)	58.5 (0.01)		66.5 (0.06)	59.3 (0.02)		71.1 (0.06)	57.6 (0.02)	
Household size, $\bar{x}$	2.5 (0.15)	2.8 (0.04)	$0.04^*$	2.9 (0.23)	2.9 (0.04)	0.92	2.0 (0.14)	2.8 (0.05)	<0.001**
Rural/urban, %			0.56			0.12			0.74
Urban/suburban	59.2 (0.06)	56.4 (0.04)		58.3 (0.07)	55.3 (0.04)		60.4 (0.07)	57.8 (0.05)	
Small city/town	32.9 (0.06)	32.8 (0.04)		37.2 (0.07)	33.4 (0.04)		27.3 (0.07)	32.1 (0.07)	
Rural	7.9 (0.03)	10.8 (0.02)		4.5 (0.04)	11.3 (0.02)		12.3 (0.05)	10.1 (0.02)	
Military Veteran, %	9.2 (0.02)	10.3 (0.01)	09.0	6.4(0.03)	2.0 (0.00)	0.12	12.6 (0.03)	20.2 (0.01)	0.05
Political Views, %			<0.001**			$0.03^{*}$			<0.001**
Conservative	23.2 (0.04)	33.2 (0.01)		21.7 (0.04)	31.3 (0.01)		25.0 (0.06)	35.4 (0.01)	
Moderate	31.1 (0.04)	39.6 (0.01)		36.3 (0.05)	40.8 (0.01)		24.5 (0.06)	38.2 (0.01)	
Liberal	45.8 (0.04)	27.2 (0.01)		42.0 (0.05)	27.9 (0.01)		50.5 (0.06)	26.4 (0.01)	
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Note. N = 4,554. Weighted estimates (Est) and standard errors (SE) shown. Statistical significance evaluated by adjusted Wald chi-square test or  $\mathbf{p} \le 0.05$ . \*\*  $\mathbf{p} \le 0.001$ 

## Gun safety policy beliefs and gun ownership

**Table 3.2** reports the prevalence of gun safety beliefs and gun ownership by gender and sexual orientation, highlighting unadjusted group differences.

**Table 3.3** reports partial results from the three multivariable models. Sexual minority men reported over three times the odds (aOR = 3.24, 95% CI: 1.79 - 5.88) of favoring a gun safety law compared to heterosexual men, after adjusting for confounding. As well, sexual minority men demonstrated about one-third the odds (aOR = 0.34, 95% CI: 0.18 - 0.65) of reporting the presence of a gun in the household as compared to heterosexual men. Gun ownership among men did not differ significantly by sexual orientation. In contrast, among women, gun safety beliefs and reporting a gun in the household did not differ significantly by sexual orientation, although the reported presence of a household gun was somewhat higher among heterosexual (31.5%) than sexual minority (23.6%) women. But, among those women who reported gun(s) in the household, sexual minority women demonstrated nearly four times the odds (aOR = 3.97, 95% CI: 1.43 - 11.03) of being the gun(s) owner as compared to heterosexual women.

I also conducted a post-hoc sensitivity analysis to assess sexual orientation differences in personal gun ownership in the population (not shown in tables). To do this, I created a binary variable in the total analytic sample (N=4,554) that identified personal gun owners (respondents who indicated that they were the owner of the gun in the household) and non-owners (all others). Gender-stratified analyses demonstrated similar patterns of personal gun ownership as in the total sample. Sexual minority men demonstrated less than one-third the odds of being a gun owner as compared to heterosexual men (aOR = 0.30, 95% CI: 0.17 - 0.56). Conversely, sexual minority women demonstrated moderately elevated odds of being a gun owner as compared to heterosexual women (aOR = 1.89, 95% CI: 0.99 - 3.60).

Table 3.2. Household guns and beliefs about gun safety among respondents in the 2010-2016 General Social Survey, by gender and sexual orientation

	T	Total		Wo	Women	·	M	Men	
	Sexual			Sexual			Sexual		
	Minority	Heterosexual		Minority	Heterosexual		Minority	Heterosexual	
	(n=195)	(n=4359)		(n=102)	(n=2401)		(n=93)	(n=1958)	
Characteristics	Wt % (SE)	Wt % (SE)	Ь	Wt % (SE)	Wt % (SE)	Ь	Wt % (SE)	Wt % (SE)	Ь
Gun safety beliefs			<0.01*			0.87			<0.001**
Favors law requiring									
police permit									
before gun									
purchase	82.3 (0.03)	73.6 (0.01)		78.6 (0.04)	77.8 (0.01)		86.8 (0.04)	68.5 (0.02)	
Household guns									
Gun(s) in household	21.5 (0.04)	35.7 (0.02)	<0.001**	23.6 (0.05)	31.5 (0.02)	0.12	18.8 (0.05)	40.6 (0.03)	<0.001**
Respondent is									
gun(s) owner <sup>a</sup>	70.4 (0.07)	65.2(0.01)	0.49		61.8 (0.10) 35.8 (0.02)	$0.04^*$	$0.04^*$ 83.8 (0.10) 92.2 (0.01)	92.2 (0.01)	0.42
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Note. N=4,554. Weighted percents (Wt  $\frac{2}{3}$ ) and standard errors (SE) shown. Statistical significance evaluated by Wald chi-square tests.  $^{a}\mathbf{n} = 1,523$  (respondents reporting presence of gun in household)  $^{*}\mathbf{p} \leq 0.05$  \*\* $\mathbf{p} \leq 0.001$ 

Table 3.3. Partial results of multiple logistic regression models evaluating gun safety policy beliefs, presence of gun in household, and gun ownership by sexual orientation and gender, General Social Survey 2010-2016

		Favors Gun Safety Law		Reports Gun(s) in Household	Ó	Owner of Gun(s) in Household <sup>a</sup>
	u	aOR (95% CI)	u	aOR (95% CI)	u	aOR (95% CI)
Overall Sample						
Sexual Minority $(n = 195)$	160	1.77 (1.14 - 2.78)*	35	0.55 (0.36 - 0.83)**	26	$3.05 (1.00 - 9.33)^*$
Heterosexual $(n = 4,359)$	3203	ref	1494	ref	1032	ref
Women only						
Sexual Minority $(n = 102)$	80	1.09 (0.62 - 1.92)	20	0.76 (0.43 - 1.35)	13	3.97 (1.43 - 11.03)**
Heterosexual $(n = 2,401)$	1880	ref	694	ref	283	ref
Men only						
Sexual Minority $(n = 93)$	80	3.24 (1.79 - 5.88)**	15	$0.34 (0.18 - 0.65)^{**}$	13	0.37 (0.05 - 2.67)
Heterosexual $(n = 1,958)$	1323	ref	800	ref	749	ref

Note. N = 4.554. Sexual minority includes individuals who self-identify as lesbian, gay, or bisexual. aOR = adjusted odds ratio; CI = confidenceinterval; Ref = referent. Multivariable analyses weighted and adjusted for gender (overall sample), age, race, education, household size, urbanicity, veteran status, political views, and survey year.

 $<sup>^{</sup>a}$ N = 1,523. Sample limited to individuals reporting presence of a gun in the household.  $^{*}$ **p**  $\leq 0.05$ ,  $^{**}$ **p**  $\leq 0.01$ 

### **Discussion**

Results from this study are consistent with a prior study conducted in a sample of U.S. veterans which found that same-sex partnered individuals were significantly less likely to report having a gun in the household as compared to opposite-sex partnered individuals. Here, I expanded on these findings to show that sexual minority women were significantly more likely to report personal gun ownership than did heterosexual women. Individuals who personally own guns report gun ownership with greater accuracy than non-gun owners who live in a household with guns; thus, injury prevention experts have called for research focused on personal gun ownership. Here, I

Elevated gun ownership among sexual minority women as compared to heterosexual women might be partially explained by the populations' disproportionate risk of stigma-related victimization. <sup>76</sup> Perceived risk of victimization is associated with obtaining a gun for self-protection. <sup>199</sup> However, sexual minority men in our sample demonstrated reduced risk of gun ownership as compared to heterosexual men despite the fact that sexual minority men also experience elevated risk of victimization. <sup>76</sup> These results highlight the importance of recognizing gender differences in the sexual minority population.

Gender norms may be one alternate explanation for personal gun ownership among sexual minority women: compared to heterosexual women, sexual minority women are more likely to report self-ascribed masculinity and traditionally male-associated hobbies and occupational preferences. Indeed, concordant with other studies, sexual minority women in our sample were more likely to report military veteran status, a masculine-associated occupation, than heterosexual women. Across studies, masculine-linked gender norms have been associated with specific patterns of health risk behaviors, including higher rates of gun ownership. As well,

this result is consistent with previous observations that sexual minority women may evidence patterns of behavioral health risks more similar to that of heterosexual men than of heterosexual women, including elevated tobacco and alcohol use, overweight and obesity, and lower rates of healthcare utilization. 90,187,188,202,203 These results underscore the public health imperative of gender-stratifying behavioral risk research among sexual minority populations, and call for additional research on the ways in which sex differences among sexual minorities shape health risks and inform tailored interventions.

Physicians tend to underestimate gun ownership among their patients,<sup>204</sup> but gun storage counseling by physicians is effective in improving patient gun storage practices.<sup>205</sup> Personal gun owners are more likely to be responsible for gun storage and safety practices.<sup>206</sup> Given that sexual minorities report favorable gun safety policy beliefs and generally liberal political views, an elevated propensity for personal gun ownership among sexual minority women could be unrecognized by physicians. Thus, clinicians working with sexual minority women should be cognizant of the populations' elevated propensity for personal gun ownership given that sexual minority women experience disproportionate risk for violence in the home, including suicide attempt <sup>6</sup> and intimate partner violence.<sup>207</sup>

The finding that sexual minority status was associated with supporting gun safety laws was not unexpected; sexual minority individuals tend to espouse political liberalism, as shown in the current sample. But, even after adjustment for political liberalism, sexual minority status was still highly predictive of favoring a gun safety law, suggesting that sexual minority populations endorse gun safety regardless of political views. However, this finding also invites several other questions regarding sexual minority gun owners. For instance, because sexual minorities favor gun safety laws, do they have safer gun ownership practices than heterosexual individuals (e.g.,

keeping a gun locked, storing ammunition separate from the gun)? This may have relevance to gun-related violence risk reduction efforts and warrants further study.

Three limitations should be considered in contextualizing these results. First, I aggregated lesbian, gay, and bisexual individuals into a single sexual minority group to increase statistical power. Relatedly, due to the relatively small sample size, the frequencies of the subgroup outcomes are small, which reduced power to detect differences. Second, 2.3% of respondents refused to answer the initial question pertaining to guns in the household; elsewhere research finds that 'refusals' are more likely to be gun owners. In analyses not shown I evaluated the impact of this effect on study outcomes and found no evidence that the refusal rate could have impacted the conclusions. Third, the measures of gun safety attitudes in the GSS were limited to a single question. In addition, gun safety practices within gun-owning households were not assessed. Additional studies with both larger sample sizes and more thorough assessment of gun-related safety concerns can resolve these concerns.

### **Public Health Implications**

Understanding accessibility to guns is a critical component of suicide prevention efforts, especially in the U.S., which has the highest rate of gun-related mortality when compared to other high-income countries. Findings from this study suggest sexual minority women may be more likely to personally own guns, a readily available lethal method for suicide, than their heterosexual peers. Given the generally liberal political views of the sexual minority population, this propensity for gun ownership may be an unrecognized risk factor for suicide. As such, I recommend that mental health providers include gun ownership questions in regular screening and counseling materials for sexual minority women to identify individuals at elevated vulnerability for suicide in the presence of other known risk factors, such as loss or depression. <sup>6,7</sup>

The following chapter of this dissertation (Chapter 4) expands upon the current findings to examine sexual orientation differences in lethal means used in suicide. In addition to highlighting access to guns among sexual minorities, these results underscore the public health imperative of gender-stratifying behavioral risk research among sexual minority populations.

# Chapter 4

Sexual orientation differences in lethal method used in suicide: Findings from 59,075 suicides in the National Violent Death Reporting System

# Background

Sexual minorities (i.e., lesbians, gay men, bisexual people) demonstrate elevated risk of suicide attempt and mortality as compared to similar heterosexuals. <sup>6,98-100</sup> But because research on antecedents of suicide among sexual minorities has generally been limited to studies assessing the risk of suicide attempt and/or ideation, rather than suicide mortality, <sup>6,210</sup> little is known about critical antecedents of suicide mortality among sexual minorities, including proximate contexts of death such as the method used suicide (e.g., firearm, hanging). This represents a critical gap in the literature given the robust association between chosen method of suicide and lethality of a suicide attempt. <sup>211</sup>

The most proximal factor associated with suicide mortality is the method used in a suicide attempt. <sup>23,210</sup> It can thus be considered the "final gatekeeper" of a suicide attempt, as it is only through the lethal method that death will occur. Case fatality rates of suicide attempt differ greatly depending on the type of lethal method used. For instance, one population-level study investigating case fatality rates of suicide attempt by method demonstrated that individuals who attempted suicide with a firearm were 55 times more likely to die than those who attempted suicide by drug/poison ingestion. <sup>11</sup> Lethal means used in suicide attempts not only vary by regional accessibility of the method (e.g., the disproportionately high firearms suicides in the U.S. as compared to other countries), but also vary by individual status characteristics. For example, the "gender paradox" in suicide – that women are more likely to attempt suicide, but men are more likely to die by suicide – stems partially from the fact that men are more likely than women to use a firearm in their attempt. <sup>55-57</sup> As well, age is associated with lethal method

such that older age has been shown to predict using a highly-lethal method (hanging, firearm, jump/fall) as compared to younger suicide attempters who are more likely to use poisoning/overdose. <sup>12,58</sup> Understanding differences in suicide method based on key status characteristics not only augments our understanding of the epidemiology of suicide but can also aid in targeting interventions to groups disproportionately vulnerable to death by suicide.

Sexual orientation differences in suicide method are understudied but might represent an opportunity for intervention. Only recently have any studies investigated this aspect of suicidality among sexual minorities. One study from 2019 investigating suicide risk in the National Violent Death Reporting System (NVDRS) among youth and young adults aged 12-29 found that sexual minority males evidenced reduced risk of using a firearm as compared to heterosexual males;<sup>212</sup> however, the analytic approach of this study has been called into question due to the large amount of missing data. 213 Another 2019 Centers for Disease Control and Prevention (CDC) study also utilizing the NVDRS compared gay male and lesbians to non-LG (lesbian or gay) decedents, and found that LG decedents were less likely to use firearms and more likely to use hanging than non-LG decedents; however this study, too, suffered from analytic issues, including considering all decedents – those classified as heterosexual and those with unknown sexual orientation – as non-LGB. <sup>214</sup> Drawing from the homicide literature, one study investigating sexual orientation differences in method of intimate partner homicide found that the percentage of homicides using a firearm was significantly higher among heterosexuals as compared to sexual minorities, but that stabbing was higher among sexual minorities.<sup>215</sup> Because intimate partner homicides primarily occur in the home, this study supports evidence that firearms might be less available to some sexual minorities as compared to heterosexuals.<sup>216</sup> Beyond these few studies, there is a paucity of evidence investigating sexual orientation

differences in suicide method that might inform suicide prevention efforts among a population vulnerable to suicide. Thus, the purpose of this study was to identify sexual orientation differences in lethal method drawn from a large, multi-year sample of suicides in the U.S.

#### Methods

## Data source: National Violent Death Reporting System

We used data from the National Violent Death Reporting System (NVDRS; 2003-2015). The NVDRS is a restricted-access CDC database that, as of 2018, pools data from 50 states and the District of Columbia on all decedents of violent death, including suicides and homicides among individuals age 12 years and older. Data were drawn from 2003-2015, which, as of 2015, pooled data from 32 states. Case records are abstracted by trained coders from police reports, medical examiner reports, hospital records and toxicology reports, witness statements, suicide notes, if available, and statements from decedents' family and friends. Each case record in the NVDRS dataset includes a series of close-ended codes regarding the decedent's demographic characteristics, circumstances surrounding the death, and information regarding the cause of death, including method of suicide.

### Analytic Sample and Sexual Orientation Coding

As of 2015, the NVDRS dataset includes data on 143,784 suicides among individuals 12 years and over. In 2012, the NVDRS added a code to denote decedent sexual orientation. To procure a final analytic sample, we first restricted the dataset to adults, age 18 and older, who died by suicide in the years in which sexual orientation was officially coded (n = 59,519). Decedents with missing data for method of suicide were dropped (n=416) as well as decedents with completely missing entries other than lethal method (i.e., missing all demographic variables

and narratives; n=28). Thus, the final analytic sample included 59,075 suicide decedents who ranged in age from 18 to 102 years old at death.

The sample includes 577 lesbian, gay or bisexual (LGB), 12,573 heterosexual deaths, and 45,925 decedents for which sexual orientation is unclassified. This large percent of unclassified data for sexual orientation represents the primary limitation of the NVDRS. Selection for sexual orientation (i.e., coded as heterosexual, lesbian, gay, or bisexual) requires that this information was included in the law enforcement or medical examiner report. For sexual minorities, this requires that a decedent was either "out" to his or her family or close friends, or reported their sexual orientation in a suicide note, journal, or social media message that was discovered soon after death. Analytic mechanisms to account for the large proportion of records with unclassified sexual orientation are described in detail below in the *Statistical Analysis* section.

# Study Measures

Dependent Variable: Method of Suicide

Method of suicide was abstracted from the code 'weapon type' in the NVDRS. Method of suicide was categorized into one of the following categories following the work of injury prevention experts: drug/poison ingestion, poison by fumes, suffocation/hanging, firearm, cut/pierce, fall, other. Lach of these categories is pre-coded in the NVDRS except 'Poison by Fumes'; this category was created using the International Classification of Diseases (ICD) code included in each decedent record in the NVDRS to extract individuals who died by car exhaust, carbon monoxide, or another gas. "Other" method of suicide included all decedents who were either pre-coded in the NVDRS as having died by "other" method of suicide, or who used a method of suicide used by less than or equal to 1.0% of the population in the NVDRS (i.e., motor

vehicle, other transport vehicle, intentional neglect [e.g., starving oneself], biological weapons, drowning, explosive, fire or burns, non-powder gun, and blunt instrument).

Primary Independent Variables

*Sex* is coded in the NVDRS as male or female and drawn from the death certificate, law enforcement report, or medical examiner report.

Sexual orientation in the NVDRS is coded as "lesbian", "gay", "bisexual,", or "heterosexual". Due to power constraints, lesbian, gay, and bisexual decedents were collapsed into a single "sexual minority" category.

Demographic and Circumstantial Variables

Race/ethnicity in the NVDRS is coded based on standards used by the U.S. Census Bureau in the 2000 decennial census, and categorized as: White, Black/African-American, Asian/Pacific Islander, American Indian/Alaska Native, Two or more, or Other.

*Marital status* is coded in the NVDRS into the following categories: Married/Civil Union/Domestic Partnership, Never Married, Widowed, Divorced, Separated, and Single, not otherwise specified.

*Military veteran status* in the NVDRS is drawn from the section denoting military veteran status on the U.S. Death Certificate.

*Region* is drawn from a code in the NVDRS denoting the state in which the death was processed and then collapsed into one of four Census region categories: Northeast, Midwest, South, and West.

In addition, the NVDRS includes the following variables to capture mental health history and substance use history: history of prior suicide attempt; history of suicidal thoughts; history of treatment for mental illness; current alcohol dependence or alcohol problem; and, current non-

alcohol substance use problem. These circumstantial variables are coded in the NVDRS as binary variables (0/1) to denote the presence or absence of the issue and are drawn from information provided in the law enforcement and medical examiner reports.

### Statistical Analysis

What to do with unclassified sexual orientation?

As aforementioned, perhaps the most critical limitation of the NVDRS is the large percentage of decedents with unclassified sexual orientation. In the current analytic sample, 77.7% of decedents have an unclassified sexual orientation. For analyses in which sexual orientation is of critical interest, appropriate analysis of these data is not straightforward, especially in determining a denominator or comparison group. Dropping all participants with unclassified sexual orientation leaves a biased dataset that is not representative of suicides in the U.S. Unless the decedent's sexual orientation is implied (e.g., reference to the decedent's romantic partner) or specifically stated (e.g., "decedent was gay") in either the medical examiner or law enforcement report, the individual's sexual orientation, LGB or otherwise, is left unclassified. Therefore, including only participants with coded sexual orientation in analysis leads to severe selection biases. 18

To ensure that the following analyses were as representative of the 'true' population of suicides as possible, analyses were conducted with a trichotomous 'sexual orientation' predictor variable, including those with unclassified sexual orientation as a reference group to which sexual minorities and heterosexuals were compared. In the general population, sexual minorities represent about 3.5% of the population.<sup>219</sup> Because of evidence demonstrating that sexual minorities are overrepresented in suicide,<sup>99,100</sup> it is expected that the proportion of sexual minorities in a population of suicides might be somewhat elevated as compared to in the general

population. Crude demographic differences between sexual orientation groups (see *Results*) suggest that the reference group with unclassified sexual orientation is 90% or more heterosexual, thus representing an appropriate reference group with which to compare sexual minorities and heterosexuals.

### Analytic Plan

Analyses were conducted in SAS version 9.4. Covariate data were analyzed to determine patterns of missingness. There was generally very little missing data (ranging from <0.01% missing 'age at death' to 8.40% missing 'military veteran status'). Bivariate analyses were conducted with the overall sample, and multivariable analyses were conducted with imputed data following Rubin's multiple imputation method. <sup>196</sup>

Statistical analyses proceeded in two stages: (a) descriptive statistics and bivariate analyses by sexual orientation were calculated to demonstrate unadjusted associations between sexual orientation and method of suicide; and, (b) multivariable logistic regression analyses were conducted to demonstrate the adjusted association between sexual orientation and each method of suicide. Given that sex is a strong effect modifier of the association between sexual orientation and mental health outcomes including suicidality, and is strongly correlated with method of suicide in the general population,  $^{6.55,100,220}$  analyses were conducted in the full analytic sample and in sex-stratified samples. The multivariable models adjusted for sex, age at death, race/ethnicity, geographic region, military veteran status, and mental health history including prior suicide attempt, alcohol dependency, and substance use. Results were evaluated at p < 0.05. We report adjusted risk ratios (aOR) and 95% confidence intervals (CI). This study was exempt from IRB review.

#### Results

Sample demographic characteristics, stratified by sexual orientation, are presented in **Table 4.1.** Significant differences in demographic characteristics are denoted by p-values from Wald  $X^2$  test or F tests as appropriate. Unadjusted odds ratios (ORs; not presented in tables) are presented in the text to expand on results. As expected, heterosexuals demonstrated few or negligible differences in demographic characteristics as compared to decedents with unclassified sexual orientation (e.g., unadjusted ORs: Sex OR=1.03; White Race/Ethnicity OR=0.94; History of Substance Abuse OR=1.04; all  $p \ge 0.05$ ). The primary notable demographic difference between these two groups was that heterosexuals were more likely to be married/domestic partnered than decedents with unclassified sexual orientation (unadjusted OR=2.91, p < 0.001); this is expected given that decedents were coded as heterosexual in part due to the mention of an opposite-sex spouse in the law enforcement or medical examiner reports.

Compared to decedents with unclassified sexual orientation, sexual minorities were younger, more likely to be female, more likely to be Black/African American, less likely to be married, and less likely to have served in the U.S. Armed forces. Variables capturing mental health history demonstrated that sexual minorities were significantly more likely to have experienced a host of mental health issues than decedents with unclassified sexual orientation. Unadjusted ORs (not shown in tables) demonstrated that, as compared to decedents with unclassified sexual orientation, sexual minorities demonstrated 2.56 greater odds of prior suicide attempt, 1.88 greater odds of prior suicidal thoughts or plans, 1.74 greater odds of a history of mental illness treatment, 1.35 greater odds of being alcohol dependent or having an alcohol problem, and 1.48 greater odds of having a non-alcohol substance use problem (all p < 0.001).

Table 4.1. Sample characteristics by sexual orientation, National Violent Death Reporting System (N=59,075)

(11-39,073)	Lesbian, Gay or Bisexual (n=577)	Heterosexual (n=12,573)	Unclassified Sexual Orientation ( <i>n</i> =45,925)
Demographics	n (%)	n (%)	n (%)
Sex			
Male	380 (65.9)	9794 (77.9)	35566 (77.8)
Age at Death (M, SD) <sup>a</sup>	38.6 (13.9)	48.6 (17.1)	47.4 (17.6)
Race/ethnicity <sup>b</sup>			
White	495 (86.1)	11168 (89.2)	40502 (88.8)
Black/African American	43 (7.5)	596 (4.8)	2889 (6.3)
American Indian/Alaska Native	10 (1.7)	254 (2.0)	552 (1.2)
Asian/Pacific Islander	14 (2.4)	248 (2.0)	810 (1.8)
Two or more	11 (1.9)	143 (1.1)	492 (1.1)
Other	2 (0.4)	107 (0.9)	367 (0.8)
Marital Status <sup>c</sup>			
Married/Domestic Partnered	57 (10.1)	6826 (54.7)	13325 (29.4)
Widowed	10 (1.8)	616 (4.9)	2959 (6.6)
Divorced	57 (10.1)	1962 (15.7)	11110 (24.6)
Separated	15 (2.7)	455 (3.6)	1233 (2.7)
Never Married	403 (71.3)	2346 (18.8)	15686 (34.7)
Single, Unspecified	23 (4.1)	286 (2.3)	861 (1.9)
Census Region			
Northeast	80 (13.9)	1922 (15.3)	6251 (13.6)
Midwest	114 (19.8)	3280 (26.1)	9619 (21.0)
South	210 (36.4)	3978 (31.6)	13316 (29.0)
West	173 (30.0)	3393 (27.0)	16739 (36.5)
Ever served in U.S. Armed Forces <sup>d</sup>			
Yes	40 (7.6)	2310 (21.1)	7965 (18.7)
Mental Health and Substance Abuse			
History of prior suicide attempt			
Yes	203 (35.2)	2341 (18.6)	8045 (17.5)
History of suicidal thoughts, plans or	203 (33.2)	2341 (10.0)	0043 (17.3)
attempts			
Yes	226 (39.2)	3924 (31.2)	11725 (25.5)
History of treatment for mental illness			
Yes	275 (47.7)	4082 (32.5)	15761 (34.3)
Alcohol dependence or alcohol			
problem			
Yes	117 (20.3)	2362 (18.8)	7285 (15.9)
Non-alcohol substance use problem	•	, ,	
Yes	119 (20.6)	1820 (14.5)	6861 (14.9)

**Note.** Statistical significance evaluated by Wald  $X^2$  test or F test as appropriate. All significant at p < 0.0001 except non-alcohol substance use problem which was significant at p = 0.002.

an = 59,073; bn = 58,703; cn = 58,230; dn = 54,110

Results from sex-stratified bivariate analyses assessing unadjusted associations between sexual orientation and method of suicide are presented in **Table 4.2.** Hanging was the most common method of suicide among sexual minorities, while firearm was the most common method of suicide among heterosexuals and those unclassified for sexual orientation. Unadjusted odds ratios (ORs; not shown in tables) demonstrated that heterosexual decedents did not differ significantly in death by hanging (OR=1.00, p = 0.79), but demonstrated moderately elevated odds of death by firearm (OR=1.19, p < 0.001) as compared to decedents with unclassified sexual orientation.

Sexual minorities demonstrated increased odds of death by hanging (OR=1.77, p < 0.001) and reduced odds of death by firearm (OR=0.43, p < 0.001) as compared to decedents with unclassified sexual orientation. Drug/poison ingestion (OR=1.56) and fall (OR=1.69; both p < 0.05) also demonstrated significantly elevated odds of use among sexual minority suicide decedents as compared to decedents with unclassified sexual orientation.

Sex-stratified bivariate analyses revealed that sex strongly modified the unadjusted association between sexual orientation and method of suicide (unadjusted ORs not shown in tables). Sexual minority men demonstrated significantly reduced odds of death by firearm as compared to men with unclassified sexual orientation (OR=0.30, p < 0.001). In contrast, sexual minority women did not differ significantly in the use of a firearm as compared to women with unclassified sexual orientation (OR=1.24, p = 0.15). Sexual minority women demonstrated significantly reduced odds of death by drug/poison ingestion as compared to women with unclassified sexual orientation (OR=0.54), while sexual minority men demonstrated significantly elevated odds of death by drug/poison ingestion as compared to men unclassified for sexual orientation (OR=2.74; both p < 0.001). Hanging was similarly elevated among both sexual

minority men (OR=1.86) and women (OR=1.67) as compared to decedents with uncoded sexual orientation (both p < 0.001).

Table 4.2 Method of suicide by gender and sexual orientation, National Violent Death Reporting System (N=59,075)

	Lesbian, Gay or Bisexual (n=577)	Heterosexual (n=12,573)	Unclassified Sexual Orientation (n=45,925)	
Overall Sample (N=59,075)	n (%)	n (%)	n (%)	P
Firearm	174 (30.2)	6871 (54.7)	23102 (50.3)	<.0001
Hanging/Suffocation	219 (38.0)	3247 (25.8)	11805 (25.7)	<.0001
Drug/Poison Ingestion	117 (20.3)	1436 (11.4)	6428 (14.0)	<.0001
Poisoning by Fumes	19 (3.3)	365 (2.9)	1133 (2.5)	0.013
Cut/Pierce	13 (2.3)	217 (1.7)	914 (2.0)	0.140
Fall	22 (3.8)	147 (1.2)	1055 (2.3)	<.0001
Other Suicide Method <sup>a</sup>	15 (2.4)	290 (2.3)	1488 (3.2)	<.0001
other survive mountains	- ( - )	( )	( )	
Women Only				
(n=13,335)	n (%)	n (%)	n (%)	<u>P</u>
Firearm	69 (35.0)	1004 (36.1)	3143 (30.3)	<.0001
Hanging/Suffocation	68 (34.5)	695 (25.0)	2483 (24.0)	0.002
Drug/Poison Ingestion	44 (22.3)	822 (29.6)	3592 (34.7)	<.0001
Poisoning by Fumes	4 (2.0)	77 (2.8)	251 (2.4)	0.531
Cut/Pierce	3 (1.5)	47 (1.7)	179 (1.7)	0.969
Fall	2 (1.0)	42 (1.5)	283 (2.7)	0.001
Other Suicide Method <sup>a</sup>	7 (3.3)	92 (3.3)	428 (4.1)	0.137
Men Only				
(n=45,740)	n (%)	n (%)	n (%)	P
Firearm	105 (27.6)	5867 (59.9)	19959 (56.1)	<.0001
Hanging/Suffocation	151 (39.7)	2616 (26.3)	9873 (26.9)	<.0001
Drug/Poison Ingestion	73 (19.2)	614 (6.3)	2836 (8.0)	<.0001
Poisoning by Fumes	15 (4.0)	288 (2.9)	882 (2.5)	0.010
Cut/Pierce	10 (2.6)	170 (1.7)	735 (2.1)	0.079
Fall	20 (5.3)	105 (1.1)	772 (2.2)	<.0001
Other Suicide Method <sup>a</sup>	6 (1.6)	198 (2.0)	1060 (3.0)	<.0001

**Note.** Statistical significance evaluated by Wald  $X^2$  tests. <sup>a</sup>Other Suicide Method includes: motor vehicle, other transport vehicle, intentional neglect (e.g., starving oneself), biological weapons, personal weapons (i.e., hands and fists), drowning, explosive, fire or burns, non-powder gun, blunt instrument

**Table 4.3** presents partial results from multiple logistic regression models evaluating the adjusted association between sexual orientation and method of suicide in the overall sample in sex-stratified samples. Sexual minority women demonstrated increased odds of death by firearm (aOR=1.45) and reduced odds of death by drug/poison ingestion (aOR=0.67) as compared to the reference group of female suicides unclassified for sexual orientation (all p<0.05). In contrast, sexual minority men demonstrated reduced odds of death by firearm (aOR = 0.40) and increased odds of death by hanging (aOR=1.41), drug/poison ingestion (aOR=2.25), and poisoning by fumes (aOR=1.17) as compared to the reference group of male suicides unclassified for sexual orientation (all p<0.001). In the overall sample, heterosexuals demonstrated elevated odds of death by firearm (aOR=1.15) and poisoning by fumes (aOR=1.19) and reduced odds of death by drug/poison ingestion (aOR=0.82) and fall (aOR=0.56) as compared to the reference group of suicides uncoded for sexual orientation (all p<0.01).

Table 4.3 Partial results of multiple logistic regression models evaluating method of suicide by sexual orientation and gender, National Violent Death Reporting System (N=59,075)

		Firearm		1	Hanging/Suffocation	ion	Dı	Drug/Poison Ingestion	ion	P	Poisoning by Fumes	es
	aOR	95% CI	Ь	aOR	95% CI	Ь	aOR	95% CI	Ь	aOR	95% CI	Ь
Overall Sample												
LGB	0.62	(0.51 - 0.75)	<.0001	1.36	(1.14 - 1.63)	0.001	1.30	(1.04 - 1.62)	0.022	1.52	(0.95 - 2.43)	0.077
Heterosexual	1.15	(1.10 - 1.20)	<.0001	1.02	(0.97 - 1.07)	0.519	0.82	(0.77 - 0.88)	<.0001	1.19	(1.05 - 1.35)	900.0
Unclassified SO	ł	1	1	1	1	ŀ	1	1	ŀ	ŀ	ł	ł
Women Only												
LGB	1.45	(1.05 - 1.98)	0.022	1.15	(0.84 - 1.57)	0.382	0.67	(0.48 - 0.97)	0.034	1.01	(0.37 - 2.76)	0.988
Heterosexual	1.26	(1.14 - 1.39)	<.0001	1.02	(0.92 - 1.14)	0.664	0.84	(0.76 - 0.93)	<.001	1.30	(0.99 - 1.70)	0.062
Unclassified SO	1	ŀ	1	1	1	ŀ	1	ŀ	ŀ	:	ł	1
Men Only												
GB	0.40	(0.32 - 0.51)	<.0001	1.41	(1.14 - 1.75)	0.002	2.25	(1.72 - 2.94)	<.0001	1.74	(1.03 - 2.95)	0.039
Heterosexual	1.12	(1.06 - 1.18)	<.0001	1.01	(0.96 - 1.07)	0.597	0.81	(0.74 - 0.89)	<.0001	1.17	(1.02 - 1.35)	0.028
Unclassified SO	1	1	1	1	1	1	1	1	1	1	1	1
		Cut/pierce			Fall		Ot	Other Suicide Method <sup>a</sup>	ıoda			
	aOR	95% CI	Ь	aOR	95% CI	Ь	aOR	95% CI	Ь			
Overall Sample												
LGB	1.14	(0.66 - 2.00)	0.631	1.33	(0.86 - 2.06)	0.197	0.52	(0.30 - 0.90)	0.020			
Heterosexual	0.89	(0.77 - 1.05)	0.170	0.56	(0.47 - 0.67)	<.0001	0.77	(0.67 - 0.88)	<.0001			
Unclassified SO	1	1	1	1	1	1	1	1	1			
Women Only												
LGB	1.20	(0.38 - 3.85)	0.755	0.34	(0.08 - 1.37)	0.128	0.74	(0.34 - 1.60)	0.444			
Heterosexual	96.0	(0.68 - 1.34)	908.0	0.56	(0.40 - 0.79)	<.001	0.81	(0.64 - 1.02)	0.075			
Unclassified SO	ŀ	1	1	ŀ	1	1	ŀ	1	1			
Men Only												
GB	1.15	(0.61 - 2.18)	899.0	1.99	(1.25 - 3.17)	0.004	0.40	(0.18 - 0.91)	0.028			
Heterosexual	0.89	(0.74 - 1.06)	0.176	0.56	(0.45 - 0.69)	<.0001	9.76	(0.65 - 0.89)	<0.001			
Unclassified SO	1				-		1	-		1.1.	2000 000 000 000 000 000 000 000 000 00	

**Note.** All analyses are adjusted for sex (overall sample), age at death, race, marital status, military veteran status, region, suicide attempt history, alcohol dependency issues, and substance abuse issues. LGB = Lesbian, Gay or Bisexual. GB = Gay/Bisexual. SO = Sexual Orientation. and substance abuse issues. LGB = Lesbian, Gay or Bisexual. GB = Gay/Bisexual. SO = Sexual Orientation. and satisfication interval. **Bolding** denotes statistical significance at p < 0.05 and adjusted odds ratio. CI= confidence interval. **Bolding** denotes statistical significance at p < 0.05 and another behavior, other transport vehicle, intentional neglect (e.g., starving oneself), biological weapons, personal weapons (i.e., hands and fists), drowning, explosive, fire or burns, non-powder gun, and blunt instrument

#### **Discussion**

Findings from this study of 59,075 adult suicides in the National Violent Death Reporting System (NVDRS) demonstrated that sex and sexual orientation were critical predictors of method used in suicide. The primary finding that emerged from this analysis was that hanging is the most common method of suicide utilized among sexual minorities. As well, sexual minority men are more likely to die by drug/poisoning ingestion than other men who die by suicide, and sexual minority women are more likely to die by firearm than other women who die by suicide. This work underscores that gender differences present in the general population (e.g., men are at elevated risk of death by firearm) are sometimes vice versa among sexual minorities; thus, there is an imperative to gender-stratify behavioral risk research in this population. This study also highlights that circumspection is required when analyzing sexual orientation data in the NVDRS, including careful consideration of an appropriate reference group of which to compare sexual minority suicide decedents.<sup>217</sup>

Findings revealed that hanging is disproportionately utilized as a method of suicide among sexual minorities. Given that a majority of suicides occur in the home, and household materials that can be used to suffocate oneself are ubiquitous, legislative or community-based method restriction interventions are unlikely to be successful in reducing hanging as a primary method of suicide. 34,52,183 Because hanging is widely used as a method of suicide among sexual minorities but is a private and highly-lethal method, research employing qualitative methods can explore motivations for hanging among sexual minorities with suicidal intent. One qualitative study with 22 presumably-heterosexuals who had survived a near-lethal suicide attempt (eight with hanging) found that hanging was adopted for two primary reasons: accessibility of method and the view that the death would be "clean", quick and painless. 221 However, suicide attempt

survivors who did not choose hanging recognized that it could be "messy", slow and painful, and believed that technical knowledge was necessary to complete the hanging. As aforementioned, reducing hanging by restricting access to the method in the general population would be unreasonable given the wide availability of materials that can be used to hang oneself. However, the socio-cultural acceptability of hanging seems to be associated with its use as a primary method of suicide. For instance, the elimination of hanging as the method of judicial execution has been linked to the subsequent rise of hangings in the United Kingdom given that hanging was no longer linked to the stigma of execution. <sup>222,223</sup> Qualitative research with sexual minorities should explore socio-cultural influences that might impact hanging as an acceptable suicide method, including the portrayal of hanging in the media, <sup>222</sup> news coverage of high-profile suicides, <sup>224</sup> and the influence of social media and social networking sites on suicidal behavior including suicide method. <sup>225</sup> A more refined understanding of socio-cultural influences on method of suicide can inform targeted awareness-raising and media messaging campaigns specific to sexual minorities.

Prior research has documented that sexual minority women are more likely than heterosexual women to be personal gun owners;<sup>216</sup> the current work extends these findings to demonstrate that, among women who died by suicide, lesbians/bisexual women demonstrate increased odds of death by firearm. Public health partnerships with gun rights/gun owner groups and gun shops are critical suicide prevention strategies that might be extended to sexual minority women. Indeed, campaigns to engage gun shops in suicide prevention efforts through awareness-raising about suicide risk and reducing access to firearms among individuals at risk for suicide have demonstrated feasibility, acceptability, and effectiveness.<sup>226,227</sup> Dubbed "Gun Shop Projects," these community-public health partnerships are currently underway in more than 20 states.<sup>226</sup>

Research is needed to understand sexual minority women's association with gun owning groups that might be appropriate suicide prevention partners. For instance, Pink Pistols is a gun rights and self-defense group specific to sexual minorities with over 50 chapters in the United States and Canada. <sup>228,229</sup> Groups like Pink Pistols that engage sexual minority gun owners can be key partners in public health collaborations to prevent suicide through strategies similar to those employed in "Gun Shop Projects".

Results showed that one-in-five sexual minority men die by drug/poison ingestion as compared to just 6% of heterosexual men. Targeted public health interventions to reduce the disproportionately high tobacco use among sexual minorities can serve as a model on which to base suicide prevention efforts targeting drug/poison ingestion and overdose among sexual minority men. Indeed, tobacco-related awareness-raising and education-based interventions have successfully reduced smoking among sexual minorities through venue-specific tailored messaging, <sup>230</sup> policy/advocacy campaigns, <sup>231</sup> and awareness-raising through sexual minority-specific media and news sources. <sup>232</sup> Similar multi-pronged interventions can be applied to drugand overdose-related suicide prevention efforts among at-risk sexual minority men.

Beyond findings specific to lethal method, the current study also demonstrated that almost half of sexual minorities had been engaged in mental health treatment prior to their suicide. This finding parallels results from population-based studies demonstrating that sexual minorities are more likely to engage with mental health services than non-sexual minorities. However, despite greater mental healthcare utilization, sexual minorities still demonstrate significant disparities in adverse mental health outcomes, including suicide. Recent evidence suggests that sexual minority-affirmative mental health treatment – psychotherapy that is attuned to concerns specific to sexual minorities, including stigmatization and internalized homophobia – is

associated with reduced psychological distress in this population.<sup>234,235</sup> Future research should seek to investigate the efficacy of sexual minority-affirmative mental health treatment in reducing suicidality among sexual minority patients at-risk for suicide.

### **Limitations and Conclusions**

There are several limitations to the current findings. The primary limitation of this study stems from the coding of sexual orientation in the NVDRS and the large proportion of records for which sexual orientation is unclassified. Estimates of method of suicide among sexual minorities in the sample might not generalize to all sexual minorities at risk for suicide because decedents coded as lesbian, gay or bisexual in the NVDRS are sexual minorities whose sexual orientation was noted – and likely salient – to their death. 218 Whether or not similar patterns would emerge among the total population of sexual minority suicide decedents is unknown. To understand this would require modernization of U.S. mortality data, including adding sexual orientation to the U.S. federal death certificate and linking electronic health records and medical databases to mortality data.<sup>217</sup> A second limitation is that, during the time of data collection, not all states were represented in the dataset.<sup>236</sup> As of 2015, 32 states were included in the NVDRS. Thus, the data might not be fully representative of U.S. suicides as a whole. In 2018, the NVDRS expanded to include all 50 states and the District of Columbia. This data is forthcoming, and it will provide a nationally-representative sample of suicides in the U.S. with which to validate the current findings. Third, data included in the NVDRS is abstracted from the medical examiner and law enforcement reports and contains many limitations similar to psychological autopsy studies, including questions around reliability of the information, whether or not the information was garnered systematically, how information was deemed relevant to include, and sampling biases. 237-239 Triangulating findings from the NVDRS with results from other post-mortem data

sources (e.g., National Death Index, international mortality data)<sup>98-100,217</sup> can be one method by which to demonstrate the validity of these conclusions.

Despite these methodological limitations, this study is innovative in the use of a dataset of suicide deaths. As aforementioned, in the U.S., an individual's sexual orientation is not included as a recorded field on the federal death certificate, a rich source for mortality data; thus, drawing from the unique coding in the NVDRS represents an important addition to the sexual minority suicide research landscape. As well, the focus on method of suicide is an advancement of current suicide-related research involving sexual minorities. Findings from this study showed that the higher rate of suicide mortality among sexual minorities, as compared to heterosexuals, is likely driven by hanging, a method of suicide that to date has not been seen as vulnerable to restricted access approaches. These findings suggest that common means restriction interventions which aim to restrict firearms at the population-level, while important, might be less likely to reduce suicide among sexual minorities. Future research should utilize both qualitative and quantitative methods to investigate motivations for hanging among sexual minorities with suicidal intent, as well as focus on developing gender-specific interventions to target safe firearm ownership among sexual minority women and reduce drug/poisoning ingestion among sexual minority men. As well, sexual minority-affirmative mental health interventions should seek to reduce suicidality among sexual minorities engaged in mental health treatment.

# Chapter 5

#### **Global Conclusion**

This dissertation sought to clarify the unique role of several risk factors for suicide among sexual minorities. Findings from these studies highlighted that sexual orientation, and specifically sexual minority status, differentially impacts both distal (i.e., psychological responses to structural stigma) and proximal (i.e., access to firearms, use of lethal method in suicide) risk factors for suicide. Three key implications emerged from these findings.

First, while these studies provided incremental insight into specific risk factors for suicide among sexual minorities, the findings as a whole showed that many of these risks are not yet well-understood and some are exceptionally challenging to prevent. For example, the first study's hypothesis, drawn from currently accepted theoretical views of structural stigma – that residing in a more stigmatizing neighborhood would result in increased psychological distress among sexual minorities immediately after a same-sex marriage ban – was unsupported in the first study. There a contradictory finding challenged the central theoretical tenets of a structural stigma theory. As well, the second and third studies showed that sexual minorities – especially sexual minority men – are protected from suicide by firearm because they are less likely to have firearms in their homes. Instead, the higher rate of suicide mortality among sexual minorities, as compared to heterosexuals, is likely driven by hanging, a method of suicide that to date has not been seen as vulnerable to restricted access approaches. These findings suggest that common suicide prevention interventions which aim to restrict firearms at the population-level, while important, might be unlikely to reduce suicide among sexual minorities.

*Second,* these studies underscored the diversity of risk within the sexual minority population as related to suicide and suicidality. Indeed, in the first study, local environment influenced

psychological responses to structural stigma; in the second study, gender modified sexual orientation differences in gun ownership; and, in the third study, gender also moderated sexual orientation differences in the method used in suicide. Together, these results underscore that the sexual minority population is heterogeneous. Efforts to explore these diversities may offer critical insights that are not captured by analyses focused on sexual minorities as a whole.

Third, these studies highlighted several methodological challenges associated with researching suicide risk factors among sexual minorities. Specifically, methodological limitations included: issues surrounding the coding of sexual orientation, as well as the lack of systematic coding of sexual orientation across datasets; power constraints, including the inability to investigate within-group sexual orientation risk differences (e.g., bisexual vs. gay) due to the representation of sexual minority status at the population level; and, finally, limitations associated with secondary data analysis, including the use of datasets that are restricted in the questions collected, especially around firearms (e.g., inability to assess firearm storage practices). Indeed, these methodological limitations highlight the importance of bias circumspection when analyzing sexual orientation in collected data as well as the need to develop better methods of identifying factors related to suicide in this vulnerable population.

The underlying goal of this dissertation was to respond to critical research gaps outlined in the 2011 Institute of Medicine (IOM) Report's recommendations to advance understanding of sexual minority population health and the 2012 National Strategy for Suicide Prevention's aspiration of *Zero Suicides*. <sup>96,240</sup> Results from the research reported here can inform public health and clinical interventions to reduce suicide risk among sexual minorities.

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